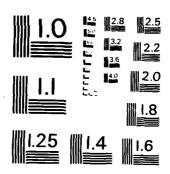
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NAVAL POSTGRADUATE SCHOOL Monterey, California



THESIS

THE DEVELOPMENT OF SELECTION STANDARDS FOR THREE NAVY RATINGS WHICH VARY IN LEVEL OF COMPLEXITY

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Kelvin W. Nesbitt

June 1983

Thesis Advisor:

Richard S. Elster

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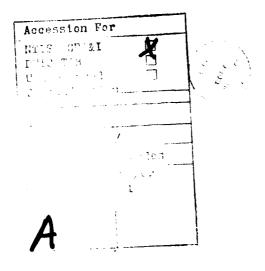
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The Development of Selection Standards for Three Navy Ratings which Vary in Level of Complexity

by

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ABSTRACT

This thesis is concerned with selection standards for three US Navy ratings which vary in terms of their complexity. The relevant literature is reviewed and a general selection standards approach is developed. approach is then applied to subsamples of a large US Navy cohort of anlistees who all had the oppurtunity of serving for at least four years. Within each rating, prediction equations are developed which link various data available prior to the beginning of the enlistee's service with three criterich measures of performance. Analyses are performed separately for groupings within ratings by race and sex. Utility analysis is employed to help determine optimum cutoffs on predictors. Many potentially useful predictive relationships are found and amongst the results is the finding that for some ratings, ability subtests are negatively related to criteria of performance. Other results are discussed and recommendations regarding implementation and future research are made.

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I. INTRODUCTION

The Armed Forces in many countries have been ardent users of the psychological techniques of selection classification. The Forces' involvement with these procedures has been for such a considerable period of time that, in fact, it could be argued that the development and refinement of these techniques has been greatly enhanced by the manpower demands of the military. In war, and in preparation for war (i.e. mobilization), the military required techniques which allowed the assessment of the training capabilities of prospective civilians for their rapid transition from civilian to uniformed status. It was inefficient to have every potential enlistee embark on a course of training when psychometric tests could be used to exclude those individuals with the lowest chances of success. Second World War, the peace time forces, largely through the development of more complex weapons systems, have some to demand much more specialized manpower skills. The Forces! requirement for selection and classification, therefore, has not only involved making military personnel out of civilians (i.e. selection), but also dealt with the correct placement of enlistees by military occupation (i.e. classification). Both the size and the diversity of the Armed Forces! manpower requirements have make it a very fertile testing ground for the application of these psychological procedures.

The present thesis is concerned with the refinement of selection standards for three employment categories (ratings) in the United States Navy. The three ratings have been chosen on the basis of the complexity of the tasks in the ratings. In increasing order of complexity they are:

Ship's Serviceman, Personnelman and Aviation Technician. All three ratings have currently applied test score entry standards based on the Armed Services Vocational Aptitude Battery (ASVAB). The aim is to examine the possibility of improving the selection criteria for each rating, taking into account rating specific utilities of correct and incorrect selection decisions. At the same time, the selection standards across jobs will be compared. Such an inter-job analysis will illustrate the relationship, if any, that job complexity has with the selection standards. This latter analysis will be qualitative (what kinds of predictors are more suited to what type of job) and quantitative (how does the level of accuracy of selection vary with job complexity).

Selection standards research for the US important for a large number of reasons. From the point of view of manpower planning, better selection standards may mean smaller training attrition, better job performance, longer careers and more rapid promotion. These outcomes have the effect of lifting the burden of maintaining the Navy's end strength from new enlistments by having greater numbers in the career force. This has even more significance with the projected decline in the services' traditional manpower pool over the next decade, particularly in that when manpower resources are scarcer, the disutility of rejecting potentially successful applicants is much greater than it is in the current recruiting climate. In a broader context, current government regulation explicitly forbids selection practices which may have adverse impacts on minority or other groups in the community (Uniform Guidelines on Employee Selection Procedures, 1978) without first investigating the use of "alternative selection procedures that are equally valid, but that produce less adverse impact" (Reilly and Chac, 1982). It is encumbent upon all employers, including the US Navy, first to determine the validity of

their current procedures and then to examine the possibility of using alternative methods which may either improve on these standards or are equally valid standards which have less adverse impact. Thus there is a strong extrinsically initiated incentive for an organization to be interested in, and to be able to measure the validity of, its own selection procedures.

There is a wealth of research on selection standards both in the US Navy and in many other employment settings. Given the aim of locking at selection, in relation to the complexity of the job and the validity of the procedures, the next chapter is devoted to literature surveys. The first section deals with the theoretical aspects of selection research, while the second is a review of the findings of previous research regarding those factors which seem most predictive of job success, in general, and for the services in particular.

The third chapter brings the information of the previous one together so as to make a clear statement of the purpose, aim, design and expected outcome of the research.

Chapter four deals with the method. First, the subject ratings are described. This includes a description of each rating, the standards currently applied and an examination of the paths through which individuals enter the ratings. It will include the complexity grading of each rating. Next comes a description of the data base for this research: where it comes from, what is included and an assessment of the quality of the data. In the next section, the personnel are described. This is followed by a section which concerns the criterion measures and predictors: What are they?, and, How they are measured? Another section describes the estimates of the utility of correct predictions and the disutilities associated with misclassifications. Next, the cross-validation procedures used for the estimation of the

validity coefficient are described. Finally, the statistical techniques are discussed.

Chapter five concerns results. The first section deals with descriptive statistics for each of the three ratings, which includes age distributions, racial and sex mixes as well as prior education levels. The next section concerns the results from using regression techniques. The final section presents results in which the direct effects of selection cut-offs are examined with respect to maximising total expected return for the selected procedure.

Chapter six is for the discussion of the results. It is divided into two sections: expected outcomes and other findings. Particular attention is paid to across rating comparisons.

Chapter seven provides a summary and a statement of the conclusions and recommendations.

II. LITERATURE REVIEW

The purpose of this chapter is to review some of the literature relevant to selection standards research. first section deals with a general overview of the technical and psychological aspects: How should selection standards be established? What is the concept of validity and how does it apply? What types of experimental design are possible and appropriate in this type of research? Then, in more detail specific theoretical issues are addressed. These are job analysis, criterion selection, predictor selection, validation and cross validation. The second section reviews research conducted specifically in selection and/or attrition. These questions are addressed: What variables have been found to be predictive of success in the work environ-Specifically, what variables are associated with success in a service environment? And finally, what factors predict success in a Navy environment? The final section provides a summary of the theory of, and the previous research in, selection.

A. THEORETICAL LITERATURE

1. Classical Model

A classical model for establishing selection standards is suggested by Drenth (1971). Like most classical approaches it is probably never rigidly applied, however Drenth's suggestion gives a good description of the steps involved and is a useful starting point for this section. These steps are summarised in Table I.

TABLE I Classical Model of Selection

Step 1. Step 2. Step 3.	Job Analysis Choice of a Criterion Try-cut Series of Tests	
Step #		~~~~
3145 A.	Testing of Experimental	Group
Step 4. Step 5.	Validation of Tests	•
STEE ST		
Step 6. Step 7.	Cross Validation	
C+ Ch 7	Tests in Final Form	
STED /	TESTS IN PIREL POUR	

Scurce: Drenth, P "Theory and Methods of Selection", 1971.

The first task is a job analysis in order that the psychological requirements of the job can be understood. In this step it is necessary to convert the description of the job into specific factors or variables which are defined and, hopefully, are measurable. The second task is to determine the criterion which is the operational means of measuring an employee's job performance. A knowledge of the criterion value should permit a ready differentiation of the good performers from the poorer ones. Ideally, it will be a metric or ordinal scale which will indicate ranges or grades of performance. In the third step, the job analysis is used to make an 'educated' guess as to what kinds of tests or other factors are important for predicting success in the job for which the standards are designed. The preceding steps are necessary precursors to the next group which deal with the empirical portion of the process. The fourth step consists of the try-cut (or pilot) tests which are administered to a representative sample of the population for which the standards are to be developed. Scores or results are recorded and retained while the employees are given a period of time to perform on-the-job. In step five, these subjects are measured on the performance criteria developed in step two, and the results are correlated with the test scores gathered in the previous step. From this analysis the merit of each test is determined leading to the a choice of tests most useful for selection and their appropriate cut cff scores. Step six is critical: the cross validation of the selection standards for new group of subjects. In this step the same procedure as applied in steps three to five is repeated using the chosen selection tests. Step seven is largely administrative. The selection tests are set up in a standard form. For small organisations this step may not be essential, but for larger ones the standardization of the tests is highly desirable, especially where there may be several different testing stations.

The Drenth classical model is a very simplistic scheme of the procedure for establishing selection standards. For one thing, it appears to assume that the organization has no selection procedures in existence, simply allowing individuals to drift into occupations. While it may not be strictly formalised, the organization probably has a good notion of what is required in the job and applies this information as a screening device. Despite these limitations, the model alludes to some very important issues which are discussed in the subsequent paragraphs.

2. Validity and Research Design

Validity is a very important issue. In the simple case above, it is described as some means of comparing the selection test results with subsequent performance in order to validate selection standards. Validity answers the guestion: do these test scores tell us a sufficient amount about the individuals' subsequent performance so that, based only on a knowledge of their selection test results, one is able

to choose the better performers. Statistically, it is usually repesented by a correlation coefficient which expresses the degree of association between the selection test scores and the measure of subsequent performance.

Two main types of research designs have been defined with respect to selection validation. These are predictive and concurrent validity (see Tiffin and McCormack, 1970). Predictive validity is that implied in the model above: selection tests are given to a group of applicants and then at some later date, criterion measures of performance are The test scores and criterion measures are then compared. For concurrent validity design, the tests scores and criterich measures are gathered virtually at the same For both types, the tests, or a subset of the tests, are considered valid, if the correlations are large enough. Until recently, the predictive validity design was considered far superior than concurrent validity approach (Drenth, 1971). However Barrett et al (1981) persuasively argue that the avidence does not support this position. They suggest that a large number of independent studies, including a recent meta analysis of 537 separate studies, have indicated virtually no difference between the magnitude of the validities found for concurrent and predictive designs. They also cite one researcher who specifically compared the designs in a national validation of selection procedures for male transit bus drivers and found that the ccefficients comparable. Mela These arguments convincing but what are the significant differences between the two validation designs?

Guich and Cranny (1982) suggest that the key difference is that for predictive designs there is a lapse of time between test and criterion measurement whereas in concurrent designs the measurements are made simultaneously. However, there is a second fundamental difference between the two:

predictive designs invariably test groups of applicants, rather than employees, as tested in concurrent designs. There are some problems for concurrent designs in relation to the restriction of range of tests scores. The employees who are the subject of concurrent designs have been preselected contributing to a narrowing of the range of data. As is noted by Hammer and Landau (1981, p. 576), range restriction shrinks the feasible range of the correlation coefficient; that is, its feasible maximum value becomes less than one. Predictive designs are not completely free from correlation problems, since there is a time difference between tests and criteria. Scores on the latter may well be due to factors other than those implied by the tests scores. Intervening variables (e.g. training factors, motivational could also inhibit criterion performances. Additionally, attrition may reduce the size of the group so that there are fewer individuals available for criterion measurement than were orginially tested. It can be seen that both concurrent and predictive designs are susceptible to shrinkage of validity coefficients.

while it seems only logical for an organization, particularly a large one, to have valid selection procedures, government regulation has a significant impact on these procedures requiring employers to have selection systems which have been tested through a general validity model. Hsu (1982) points out that the Uniform Guidlines on Employee Selection Procedures (1978) specify "that a necessary condition for use of a selection procedure, when that procedure produces adverse impacts on the hiring rates of certain minority groups, is that this procedure be demonstrated valid" (Hsu 1982, p. 509), but that this is not a sufficient condition for its use. Hsu explains that for an organisation to use any existing selection procedure that produces adverse minority impacts, it must have higher

validity than alternatives which have no adverse impacts. Many civil actions have been won during the last decade by employees (see Rowland et al. 1980, pp. 64-85) against employers who have had selection systems for which the validity has been unknown or inadequately established. So there is a considerable amount of external pressure for the organization to, first, know what the validity of its procedure is, and to, second, establish, in the case that there are adverse impacts, that there is no alternative system which is equally valid that will produce a less adverse impact.

Given this general overview of the technical literature, it is now possible to deal in more detail with other issues.

3. Job Analysis

It is probably a fairly safe assumption that very faw attempts at introducing selection standards begin with a thorough job analysis (such as is described in Dunnette, 1976, Chapter 15). This is so for a variety of reasons, not the least of which is that it is very time consuming and requires a considerable amount of effort with, at best, very little by way of quickly identifiable return for these labours. A job analysis is not part of this research, but Chapter IV includes a brief job description of the three Navy ratings involved, in order to put them into an appropriate context. Presumably thorough job analyses have been done for these ratings and used for the development of the selection procedures.

4. Criterion Selection

Dunnette's bcok ("The Handbook of Industrial and Organizational Psychology", 1976) has two excellent chapters on the selection of an appropriate criterion, which is the

second step in Table I. This is perhaps the hardest step of all and has been one of the "key problems in industrial and organizational psychology, as evidenced by the massive efforts designed to clarify its theory and improve its measurement" (Smith, 1976, p. 745). Smith specifies three requirements of a criterion of performance:

- a. that it be relevant of some organizational or individual goal, but neither biased nor trivial;
- b. that it be reliable and "involves agreement between different evaluations, at different periods of time and with different although apparently similar measures" (Smith, p. 746);
- c. that it be practical: that is, available, plausible and acceptable to those who will use it for
 decisions.

Another issue that has received much research is the choice between a single performance criterion and multiple criteria. In the classical model, the emphasis was on a single criterion which met all the requirements mentioned above (Guion, 1976, p. 783), but the weight of recent opinion is towards multiple criteria (Smith, 1976, p. 747). The first argument in favour of the multiple criteria solution is on logical grounds: the various possible criteria are all different ways of measuring the same unitary concept and using more than one measure gives a tetter estimate of this concept's true value. It could be that two individuals with equivalent total performances could achieve these results through quite different behaviour patterns. The other arguement for multiple criteria is empirical: "an overwhelming majority of studies involving statistical analyses of sets of criterion measures finds that these analyses rarely yield a single general

factor...several criterion measures are necessary to account for the variance in a criterion correlation matrix" (Smith, 1976, p. 748). The choice of criteria and how to combine them are difficult processes and must all too often be based on convenience rather than their relevance to long term effectiveness.

Several recent studies have proposed the use of utility concepts in performance measurement which could be used as a criterion (Iandy, Farr and Jacobs, 1982; Schmidt and Hunter, 1981). Utility concepts include such things as the number of successful employees hired, increases in the average level of performance, and the amount of savings in dollars and cents, as a means of evaluating the effects of various selection strategies. Most recent applications of utility concepts deal with the value to the organisation of the hired individual which was derived by estimating the average value and the standard deviation of performance expressed in dollars. These criteria as they stand are inappropriate for the present study, but the application of utility analysis is discussed below in relation to the determination of optimum cutting scores on predictors.

In addition, Smith (1976) suggests several other variables which have been found useful as providing "hard" and "soft" criterion measures. Under the hard category are tardiness, absences, accidents, tenure or turnover, sales, production, job level and promotions. Even these may contain some soft elements largely due to the effects of human judgements. The soft criteria are those which involve the use of rating scales, usually completed by supervisors. There are a large number of sources of rating scale error, and an even larger number of methods to minimise their effects (see Landy and Farr, 1979; Smith, 1976, pp. 757-764), but these will not be considered in detail here.

5. Fredictor Selection

In trying to select tests which may serve as predictors, Guion (1976) suggests three criteria:

- a. complex behaviour (i.e. the criterion) cannot be fully predicted by simple means, thus a single predictor is usually inappropriate:
- b. complex performance is a function of the individual, but only to a certain extent, therefore in thinking of predictors, it is pertinent to go beyond individual traits and include situational and demographic variables as sources of potential predictors; and
- c. complex behavior is not likely to be optimally predicted in precisely the same way for all people, while it is not feasible to have prediction equations for each individual, it may be possible to develop different equations for logically different subgroups.

He also suggests many publications to consult (e.g. Ghiselli, 1966) which provide a summary of predictive validities for various job categories, by type of predictor and by type of criterion. They are, he suggests, a good starting toint.

6. Validaticn

To this point the concept of validity has been described as a correlation coefficient which expresses the relationship between predictor and criterion measures. However, this concept can be extended. As noted by Campbell (1976), the correlation coefficient approach to validity implies:

a. "there is one normally distributed criterion or criterion composite;

- b. there is one normally distributed predictor or predictor composite: (and)
- c. the relationship between them is linear and homoscedastic." (p. 205)

These are stringent conditions and are often violated. He further challenges the practical use of a validity coefficient: how are decision makers to use say, the value of .55, in determining which of a group of applicants should be selected? The coefficient says little about how to produce the desired outcome. Provided the above three conditions are met, the validity coefficient is useful for determining which of a group of predictors is more appropriate for selection but Campbell describes what he calls "decision centered" validity as a more useful approach.

With a decision orientation, selection research is viewed an attempt to predict discrete cutcomes on a criterion which is categorical in nature. The categories to be predicted are made up of those who succeed against those who fail. The task for the predictor then is to establish a "cutting" score which maximises the proportion of correct predictions. But how is the proportion of correct predictions to be defined? In the equations below, the figures A, B, C, and D all refer to the numbers of personnel by predictor/ criterion cutcome. A is the number who are predicted to pass the criterion and who do, in fact, pass. the number who are correctly predicted to criterion. C is the number who are predicted to pass but who subsequently fail the criterion. Finally, D is the number who are predicted to fail but who would have passed. these figures, two examples of how the proportion of correct classifications may be defined are as follows. One method, Equation 2.1, is to take the proportion of correct classifications over the total number to classify. Another

approach, Equation 2.2, takes as the statistic the number of successes over the total number who pass. There could be other possible definitions. However for both equations, the idea is to select a "cutting" score on the predictor which maximizes the value of the expression. What is forgotten in both these definitions is that each embodies a value judgement in the way the statistic is defined. Equation 2.1

$$CP = (A + E)/(A + E + C + D)$$
 (egn 2.1)

$$CP = (\lambda)/(\lambda + C)$$
 (eqn 2.2)

applies an equal weight to both correct prediction outcomes, and does not distinguish between the two types of error (categories C and D in the equations) in terms of their (negative) value to the organisation. Equation 2.2 ignores completely the outcomes for individuals who are not in the predicted success group, in effect assigning a value of zero to each of these outcomes. The question arises then as to what values should be assigned to each outcome for the four possibilities, given in the equations, to derive an appropriate outting score?

The answer lies in the application of the concept of utility which was referred to briefly above. As described by Campbell (1976), utilities are estimated for each outcome in Equations 2.1 and 2.2. For simplicity in the following discussion, the terms pass and fail refer to criterion results whereas positive and negative refer to predictor score results. Thus, outcome A is a pass/positive outcome: these members are correctly predicted to pass on the criterion. An example of the application of utilities and how they are maximised by varying the cutting score is given

TABLE II

Total Expected Utility for a Predictor Cut-off Value

A. Fixed relative utilities for each outcome

Success -5 10
Failure 10 -30
Predict Predict Failure Success

E. Probability of each outcome for a certain cut-off

Success

.16
.24
.40
Failure

.48
.12
.60

Predict
Failure

Success

C. Freducts of utilities and probabilities

Failure

-.80
2.4

Failure

Predict
Pailure

Success

Total expected utility for this cut-off = 2.80

in Table II (this example is very similar to that given by Campbell, 1976, p. 212). As can be seen from Table II, part A, the pass/positive outcome is given a value of 10 units of utility which is the same value as the fail/negative outcome. This implies that it is as valuable to the organisation to predict a success as it is to correctly predict a failure. The negative signs in the other two cells suggest that each outcome is a disutility to the organisation. Further, it is clear that the relative values in these cells indicate that the disutility of incorrectly identifying a

failure, is far greater than incorrectly identifying an individual who would have been successful. In fact the former disutility is six times as great and it is not difficult to see why this would be so. Allowing a future failing individual in to the organisation is much more expensive because of the larger amount of investment which is lost (e.g. training expenses, salary) compared with the relatively small investment lost on the rejected applicant who would have been a success. The question remains however as to how to use these utilities in the deriving selection out-offs.

To see how the utility values are used for determining the effectiveness of a selection procedure. necessary to examine all three parts of Table II. Part B shows the proportion of the sample residing in each cell based upon a cut-off score on the predictor. It forecasts that 38% of the applicants will be successful. Treating these cell entries as a probability of belonging to that cell, it is now possible to calculate the expected utility for each cell by multiplying the probability by the corresponding utility. The result is shown in part C of the table. For a cutting point of 38% predicted successes, the total expected utility is simply the sum of each of these cells: 2.9 is the expected utility for this particular cutting Using this same procedure the next possible cut-off score on the predictor can be evaluated. Clearly, this cut-off will most probably produce different proportions of cases in each cell, which will in turn change the value of the total expected utility. Continuing this process for all possible cut-offs produces a total expected utility for each cut-off level. The cut-off which should be chosen is the one that maximises the expected utility. Of course if all cut-off levels result in negative expected total utility, then the selection device is of no use to the organisation and it should not be used.

Within the US Navy setting, it is possible to make estimations of the various utilities associated with each of the four possible predictor/criterion cutcomes through cost and productivity analyses by rating. Utilities for three of the four possible outcomes appear to be relatively straightforward, provided some simplifying assumptions are made. These three (associated with A, B and C in the equations) may be estimated from the Billet Cost Model (Assessment Group, 1983). However, the disutility of the other cell, misclassified pass (D in the equations) is more difficult. For the defense force the disutility of this outcome is very much a function of the recruiting market. During lean recruiting times, when quotas are hard to fill, to reject a potentially successful candidate has greater disutiltiy than it would in the current recruiting climate. The utilities associated with the other three cells should be impervious to such fluctuations. The actual derivation of utilities is discussed in the Method chapter.

7. Cross Validation

Depending on the statistical technique employed, the results of selection standards research will be effected by certain characteristics specific to the sample used in the study (Weiss, 1976, p. 332). This means that the results will reflect not only the true relationships between the criterion and predictor variables, but they will also reflect relationships which are unique to that particular sample, and this uniqueness may not be generalised to the population. To avoid these problems selection results are usually cross-validated.

It is not necessary to cross-validate in the same manner which was described for the classical selection model. This would involve the complete replication of the original research on a new group of subjects. More efficient

and less costly methods are available. Campbell (1976, p. 214) describes what he calls the 'double cross-validation' design. Here the sample is randomly split into two equally sized subsamples. The statistical techniques are applied independently to each subsample to derive the decision rules for the final version of the selection device. The decision rules for subsample 1 are applied to subsample 2 and the validity estimated. The decision rules for subsample 2 are applied to subsample 1 and another coefficient is calcu-The average of these two coefficients gives the validity for the procedures. Ideally, of course, the decision rules for each subsample will be essentially the same. Marked deviation between subsamples indicates that either one or both results are sample specific and therefore they would be suspect.

The double cross-validation design could apply equally well to the utility method for estimating the optimal cut-off values. Cut-offs could be determined on predictors for one group and then compared with the cutoffs derived from these same predictors on the other subsample.

B. RESULIS FROM PREVIOUS RESEARCH

As was indicated earlier, there is a great deal of research literature on personnel selection within organisations. This includes studies for almost every conceiveable employment type both across the community and, within and by, the services. The purpose of this section is to include sufficient of the previous research so as to get a basis for formulating research hypotheses in the next chapter. It is neither intended nor claimed that what follows is an all inclusive review of personnel selection research.

1. General Results

The history of mental testing is intertwined with that of rersonnel selection in organisations. During the 1950's and 1960's, many organisations adopted the IQ as a standard measurement for use in selection and placement in virtually any employment setting. But the whole area of human ability testing, including the concept of IQ, came under fire in the late 1960's. The civil rights movement had a lct to do with this, as articles and monographs (e.q. Jensen, 1973) indicated that there were significant mincrity score differences on these tests. Further, there were claims that minority groups were disadvantaged, through test bias, by organisations which employed psychological tests in their selection procedures. In a recent article Carroll and Horn (1981) sought to make a distinction between the science and the application of human ability measurement. claimed that the science of human abilities was in a "hase of confusion (as evidenced) when ideas about human abilities are regarded as equivalent to ideas about IQ measurements (and) that valid criticism of particular applications of measurements are not, in general, criticisms of the science of ability measurement" (page 103). They argue that there is no particular fault with the scientific aspects of human ability measurement, claiming that 80% of the variance of ability differences is accountable in test measurement, and suggesting that the application of these tests the difficulty.

Schmidt and Hunter (1981) also discuss this issue, but in a more applied sense. They have derived, over the last five years or so, a procedure for evaluating the significance of individual validity studies, in which they say it is sometimes difficult to detect the presence or effect of sampling bias. Specifically, they directly address many of

the criticisms of mental testing put forward in the last decade. These include the notions of low utility (selection procedures have little direct bearing on organisational productivity), test unfairness (selection procedures are tiassed against minorities) and test invalidity (selection test validities may be situation or job specific and not truely general). Numerous recent studies are referenced in support of the idea that selection procedures do not have low utility. This leads to the conclusion that the potential effects of employment testing on productivity have been underestimated. The authors deal with test unfairness by suggesting that the factors causing low test scores also produce poorer job performance, for both majority and minority groups. They conclude that employment tests ic not cause adverse impacts against minorities and they seem to support Jensen (1980), who demonstrated that when minorities are selected on procedures developed on white samples the minority group is favoured (see Jensen, 1980, p. 515). third issue of test invalidity is similarly argued with the presentation of many studies which support the notion that tests are valid for employment prediction. One relevant study showed that when jobs were grouped according "to complexity of information-processing requirements, validaties of a composite of verbal and quantitative abilities for predicting on-the-job performance varied from .56 for the highest level job to .23 for the lowest" This is an interesting finding considering the emphasis in this thesis on selection for various complexity levels. It suggests, first, that cognitive ability should be valid for predicting job performance and, second, that the validity should increase with job complexity. Smith and Hunter reach three conclusions:

a. professionally developed ability tests are valid

predictors of performance both on-the-job and in training for all jobs;

- b. ability tests are fair to minority groups as they do not underestimate the expected job performance of minorities; and
- c. ability tests used in selection can effect huge labour and cost savings in an organisation.

In contrast to Schmidt and Hunter's paper, which makes a frontal attack on the critics of ability testing, Reilly and Chao (1982) examine the validity and fairness of some alternative selection procedures as is required by qovernment regulation where there are adverse minority effects. Of the eight alternatives, two are relevant for These are biographical data and academic the present study. performance measures. Biographical data is that type of information usually given by an application for a job, and includes such things as age, number of years schooling, marital status and number of dependents. These data have been successfully used in predictive cross-validated studies, returning validities not as large as for ability tests, but sufficiently close to suggest they could be a feasible alternative. Academic performance measures, as grade point averages and course results, have not fared so well. Validities based on performance criteria for these items have been either insignificant or extremely low. They are not recommended as an alternative.

In sum, therefore, it is clear that mental ability tests are relatively valid predictors of job performance in virtually all settings, and that biographical data are also useful and valid as predictors. Higher validities have been found for more complex jobs when compared with jobs of low complexity.

2. Military Research

As was indicated in the introduction, the military has been a very fertile ground for research in, and the application of, selection standards technology. Almost 150 studies on the US military were recently reviewed by Vineberg and Joyner (1982). These occured between 1952 and 1980 and were selected because they were job performance prediction studies. The authors' paper dealt with criterion and predictor variables and the usual technical measure of validity (i.e. the correlation coefficient) associated with different combinations of these variables. This extensive survey article is discussed in the next few paragraphs.

Vineberg and Joyner categorise the criterion variables used in military studies into three types: proficiency, job performance and suitability. Proficiency variables are objective type measures of either job knowledge, task and/or task element performance. Job performance is expressed as either global ratings, job element ratings (both of which are subject to the rating scale errors referred to above), productivity (applied only to Army recruiters) and the grade or skill level obtained. The suitability criteria are made up of such things as length of service, completion of enlistment term, miscenduct measures and recommendations for Within and across these subdivisions the ra-enlistment. best median validities obtained were for job knowledge (.40), task performance (.31), suitability (.24) and global rating (.15) from a total of 350 validity coefficients. Based on these values, the authors conclude that the task performance and suitability validities are high enough to suggest they would be useful in selecting military personnel.

Many different types of predictors have been used as the independent variables in selection research. The most popular is aptitude (e.g. mental tests), followed closely by biographic and demographic variables. In most aptitude studies at least one other type of predictor was used. Age and education level (expressed as number of years schooling) were the most frequently used biographical variables. When predicting suitability criteria, the variables age, mental ability and education have consistently been found to be significant. The median validities are: education (.36), mental ability (.24) and age (.21). When all three of these variables are taken into account simultaneously, then validities have ranged from .24 to .39. Some evidence is presented which suggests that when some early military experiences are taken into account (e.g. ratings at recruit training), then the validities for prediction range from the .40s to the .60s.

While the authors conclude that the research suggests the usefulness of these predictors to forecast performances on these criteria, they make two relevant points about the relationship of the variables and how the validities may be improved upon (page 22). The first point is similar to one made earlier in this paper: likely to be many important differences within and across jobs, and these need to be taken into account when theorising about what kinds of relationships exist. differ in terms of job difficulty, the level of effective cailing of performance and how many of the encumbents reach this level of performance. These kinds of factors effect what the criterion measure should be and how it should be measured. The present study hopes to address specifically the issue of job complexity.

The second point is that predictor and criterion relationships are not likely to be static over time. With experience on-the-job, those predictors which were useful for predicting "early" criteria of performance tend to "wash out". It may be that as training and experience tend to

make the differences in technical proficiency smaller and less significant, then predictors are less able to discriminate between the performances. These points tend to empasise the notion that the expected relationships between predictors and criteria might not be found in a sample, but also that there are probably many alternative explanations for insignificant results.

Studies of predicting military performance have been done in other countries, with similar kinds of results. Some of the more recent ones are summarised by Johnson An earlier paper, quite relevant to the present study, was done by O'Gorman (1972). He studied regular entrants to the Australian Army and found that a total of ten pre-enlistment variables combined in linear regression were useful in predicting an effectiveness measure of performance. The criterion consisted of assigning the value "1" to those who did not complete their service for reasons other than death or illness, those who completed three years without promotion, and "3" to those who completed three years and who were promoted to corporal or higher. The significant predictor variables included age, marital status, aptitude scores (general ability and clerical ability), results on a psychiatric inventory, hirth order and prior service.

More specifically for the current thesis, there has been some recent pilot research (Lurie 1981) about determining selection standards for the Ships Serviceman rating. In this study Lurie attempted to relate pre-entry variables to an improved criterion of Navy success: rather than the survivability of recruits during their first term he attempted to employ a job performance criterion to determine selection standards. His criterion was a combined advancement and survivalability measure, the object being to predict this variable given AFQT score, age, number of

TABLE III
Percentage of SHs scoring below designated AFQT Score

<u>Grade</u>	AFOT =< 20	AFOT = < 50
-F-1	54	83
F-2	50	85
E-3	45	7 Ř
E - 4	34	70
E-5	30	6.2

Source: Lurie, P.E. "Relating Enlistment Standards to Job Performance: A Pilot Study for Two Navy Ratings" CNA No. 81-0048 of 15 Jan 1981.

primary dependents and years of education for a cohort group of enlistees. His study used a cohort which joined the Navy in 1973, and records of their service were available up until 1977, allowing for advancement to E5. Regression models were used in calculating the relationships between the variables which allowed for the construction of "service status" protabilities. These probabilities show for various amounts of time in service the likelihood that the sailor will be either advanced/demoted to a particular rank, or discharged from the service. Without controlling for education, Lurie's data (see Table III) show an expected relationship between grade and AFQT. While the numbers are a little confusing, it is clear that as rank increases then the proportion of low AFQT scores declines. However, what is not obvious from Table III is the combined effect of high school graduation and AFQT. Lurie found that high school graduate enlistees with higher AFQT scores, have a better chance of advancement to E3 and E4, while the lower AFOT scorers have a higher rate of attrition. For non-high school graduates, the probabilities indicate that recruits with

lower AFCI scores outperform those with the higher scores. This rather striking result cannot be attributed to differential attrition rates as these are virtually identical for both low and high AFCT scorers. Non-high school graduate recruits with lower AFOTs advance more guickly to E4 and E5 than do high scorers. On the other hand, high scorers have a much higher chance of being reduced in rank than the low scorers. Lurie suggests that perhaps non-high school graduates with high AFQTs are dissatisfied with their placement as SHs believing they should be better employed. However no data is presented to support this notion. His overall conclusion is that it appears feasible to set enlistment standards for this rating based upon performance criteria. Unfortunately, his recommendations do not suggest any new standards.

C. SUMMABY

1. Theory

It is now possible to give a short summary of the theoretical literature in this chapter. This is given in point form.

- there are both intrinsic and extrinsic pressures for an organisation to be interested and able to validate its selection procedures.
- validation of selection standards can take either the 'predictive' or 'concurrent' research design as they produce equivalent results.
- all selection research should begin with a thorough job analysis so that relevant information is obtained to assist in the development of both criterion and predictor concepts.
- certain principles are known which assist in the

selection of criterion and predictor measures.

- validity coefficients are susceptible to shrinkage due to a number of factors which may be controlled or adjusted for.
- while correlation coefficients are useful in selecting the most relevant predictors from a large number of potential predictors and as a technical statistic of validity, alone they do not indicate precisely what the selection procedures should be.
- providing utilities can be estimated for the various predictor/criterion outcomes, the decision centered validity design appears the most practical and useful to determine predictor cut-offs which maximise utility and which ultimately determine the selection system's cverall usefulness.
- the double cross-validation design is a means of overcoming some sources of sample specific bias in selection standards research.

2. Empirical Results

From the review of previous studies the following points may be summarised.

- intra- and inter-job differences (e.g. complexity) are likely to effect the relationship between predictor and criterion measures.
- validity coefficients for more complex jobs are likely to be higher than less complex ones.
- relationships between criterion and predictor variables are not likely to be stable over time as training and experience appear to "wash out" these

relationships by reducing differences in criterion performances.

- significant predictor measures found in military studies include age, marital status, education, ability test scores and other biographical variables.
- criterion measures used previously in the military have included the following either singly or as composites: length of service, advancement, recommendations for re-enlistment and various misconduct measures.
- scme recent studies (e.g. Lurie, 1981) have emphasised the unusual moderating effect—some variables may have on predictors.

III. AIM AND PURPOSE OF THIS RESEARCH

This chapter is designed to make a clear statement of the aim and purpose of the thesis. It is a chance to synthesise the literature survey of the previous chapter into a form which indicates the expected cutcome from the analysis of the ratings chosen for study. The general design of the research and a statement of the research hypotheses are also given.

A. PURPOSE OF RESEARCH

The overall purpose of the present research is to analysise the available data on a cohort of entrants to three US Navy ratings in order to test the possibility that the current selection procedures can be improved. The ratings are Ships' Serviceman, Personnelman and Aviation Technician. The study involves the development of selection standards which it is hoped will enhance the total effective performance of future entrants to these ratings.

There are several secondary purposes. Inter-rating differences will be examined in relation to the level of complexity of each rating. To this end, ratings have been chosen deliberately with the intention of having a wide spread in terms of job complexity. The research is intended to examine rating differences with respect to different criteria of performance both for single and composite measures. Finally, it is intended to apply the concepts of utility analysis (that is, applying estimates of the cost and the productivity of sailors) as a means of determining the selection procedure which will yield the maximum pay-off in terms of total expected return to the Navy.

B. AIM

The aim of this thesis is to investigate the possibility of improving the selection standards for three US Navy ratings.

C. DESIGN

In terms of the theoretical literature, the design of the present study is more 'predictive' than it is 'concurrent'. It is not strictly 'predictive' in the classical sense, as it does not include data on all applicants for the US Navy, but rather has information on applicants who were subsequently enlisted. The research is somewhat 'concurrent' in that sailors within a rating are included for analysis provided they met at least one of the three criteria of inclusion discussed in a later section.

After the selection of appropriate predictor and criterion measures, the next feature of the design is the use of double cross-validation coefficients as an overall gauge of the usefulness of the relationship between predictors and criteria as a selection tool. This involves the random splitting of the total sample into two halves, developing a linear regression relationship for each half, then testing the correlation between predictions and criterion measures in opposite halves, and averaging these correlations to get an average validity coefficient. In essence, using the data from one half to predict the behaviour of the other half.

Using these linear regressions models, it is intended to apply estimates of utility and disutility (i.e. the extent to which a particular selection outcome gives a positive or negative return to the employer) in order to establish the optimal cut-off value of predictions to maximise the total expected utility. This cut-off value will be expressed as a

percentage improvement on the total utility obtainable when the predictive equation is not used.

The design also includes an analysis of the effects of moderator variables (these are variables which interact with the predictors). One of the most important in this study is job complexity. Others which have been shown to be significant in other studies, and which will be addressed in this study, are race and sex.

D. RESEARCH HYPOTHESES

While this research in more empirical than theory based, it is possible to formulate at least two research hypotheses based on past findings.

1. Hypothesis 1

Age, educational level and mental ability/aptitude will be amongst the variables found to be significant predictors of criterion performances.

2. Hypothesis 2

Validities of the predictor/criterion regression equations will tend to increase as the level of complexity of the Navy rating increases.

IV. METHOD

This chapter describes the execution of the research. On a section by section basis it deals with: the ratings selected, the data base, the variables used, the subjects chosen and the statistical analyses employed.

A. NAVY RATINGS SELECTED

As has already been mentioned, the ratings in this study are Ships Serviceman, Personnelman and Aviation Technician. They were chosen because they represent varying degrees of job complexity. According to a complexity scale (see Sands, 1979) these ratings have values of:

a. Ships Serviceman : 40

b. Personnelman : 67

c. Aviation Technician : 95

The whole scale has a range from 10 to 99, with a median value of 70, and the most complex rating assigned a 99. These ratings therefore are representative of the spread of complexity of ratings in the US Navy.

Admittedly, the ratings differ in terms of the nature of the work as well as in complexity. This could well be a significant cause of rating differences and may influence the outcome of the study. A more ideal approach would have been to select different levels of complexity within the same type of employment. However, the data did not allow such a fine classification of personnel. Another variable which will obviously have an effect on the outcome of this study is the fact that for all ratings an existing selection procedure is in use. It requires that entrants exceed particular ASVAB subtest scores, as well as reaching a certain SCREEN score. The SCREEN score is explained in detail later, but it is a probability which represents the chances of the applicant completing his period of service (Sands, 1979). The probability is derived from a consideration of race, AFQT, marital status and high school diploma status. Such prescreening will tend to restrict the range of both predictor and criterion variables which will have an effect on validity coefficients.

A general description of each rating, which is taken principally from the US Navy's "Enlisted Career Guide 1981-82", follows.

1. Ships Serviceman (SH)

At a primary level, depending upon their specialty, the SH may work in a barber shop, ship's store, laundry, dry cleaning plant or office. Subsequently they may assume more responsibility in supervisory and managerial roles associated with retail stores, commissary stores and laundry/dry Among the distinguishing characteristics cleaning plants. of this particular rating is an emphasis on basic interpersonal skills. While the actual employment may vary in physical surroundings and requirements, the jobs all have a lot to do with dealing and interacting with other people. An affinity for dealing with people is therefore essential. Although there is not a requirement for high level cognitive skills, it is nonetheless important for SHs to be of above average ability in arithmetic and in such things as recordkeeping and detail work. After recruit training, attends a 6 weeks' course in the trade. There are about 5,300 personnel amployed in this rating at present.

2. Fersonnelman (PN)

The Navy's Fersonnelman is a rating which specialises in clerical and counselling duties with respect to personnel. They provide information and counselling to enlisted sailors related to Navy occupations, opportunities for general education and job training, promotion requirements, and in rights and benefits. After recruit training, the FN completes a 6 to 7 week technical school. There are about 7,000 personnel employed in this rating at present.

3. Aviation Technician (AT)

full title of this rating is Aviation Electronics Technician. As the title suggests, sailors in the rating are employed maintaining advanced radio, radar and electronics equipment that are carried bracd no aircraft. Within this broad area of responsibility, there are three categories of employment. The AT may be involved in testing and analysis of equipment, its maintaince and repair, and in related administrative tasks. After recruit training. the AT attends up to 6 months of full-time schooling, and depending upon whether the AT is a four or six year obligator, he/she goes on to do a further 6 months of advanced first term avionics. The rating employes about 10,000 personnel at present, and there are shortages of ATs.

B. CATA BASE

The cohort data base was assembled by the Defense Manpower Data Centre (DMDC) for the Administrative Science Department of the Navy Postgraduate School (NPS). Cohort data on some 206,000 non-prior service sailors were obtained through the merging of three usually separate Navy files. These files are referred to as the DMDC file, the Advancement file, and the Navy Health and Research Center (NHRC) file.

The DMDC file contains pre-enlistment variables as well as entry dates and current status. It was assembled by linking various copies of the Enlisted Master Record across time using Social Security number. This number was deleted in the final cohort file in the interests of privacy.

The Advancement file records variables related to the promotion of individuals through the advancement system. For example, this file provided values for the Final Multiple, a composite score, which determines whether the sailor is promoted.

The NHRC file contains many of the same pre-entry variables as the DMDC file, but records additional pre- and post-entry variables. Amongst these are aptitude scores on the Armed Services Vocational Aptitude Battery (ASVAB) subtests and percentile scores on the Armed Forces Qualification Test (AFOT). The file also contains such things as the total number of absenses without leave (AWOLS), demotions, desertions and the times for promotion to various grades (eg number of days to E4). NHRC is the largest of the three files.

The cohort file included data from all enlisted entrants to the Navy between 1 September 1976 and 31 December 1978. No special data collect or monitoring of these subjects was set up; rather snapshots of their progress were taken in September 1982 from the various files to create the NPS cohort file. The time period allowed each enlistee the oppurtunity of serving about four years, with some being abla to serve as long as six years.

C. VARIABLES IN THE CATA BASE

There are over 260 individual variables in the cohort data base, and those relevant for this study are described below.

1. Criterion Variables

Three critericn variables are used. One is unitary while the others are composites. The singular criterion is total length of service in months. This is available on all members of the cohort and is a true metric scale suitable for regression analysis. The second criterion is a positively oriented ordinal scale. Three mutual exclusive and collectively exhaustive categories of criterion outcomes are identified. Those individuals who did not complete four years of service for other than medical, death or officer entrance reasons, were grouped in the first category. Those who completed four years' service were placed in the second category. Those who served four years and were rated (i.e. became qualified in the rating), graded E4 and recommended for re-enlistment were placed in the third category. the purposes of statistical analyses, these three categories were given the values "10", "20" and "30" respectively. Since this scale attempts to differentiate the better sailors from the others, it is known as the 'goodquy' scale. The third criterion is negatively oriented and is known as the 'hadguy' scale. The first category includes those who did not complete four years of service, who had been discharged for negatively criented reasons, and had either desertion, AWOL or confinements recorded against them and received a value of "10". The second ordinal category includes those who did not complete four years of service, or those who had demotions or were not recommended for re-enlistment regardless of their length of service and was valued "20". The remainder of the group, who did not exhibit any particularly negative behaviours, were placed in the third category which was valued "30".

While these criteria could be defined in different ways, they seem to be consistent with the three requirements suggested in Chapter II.

2. Predictor Variables

The following potential predictor variables are contained in the cohort file. They are:

- a. age at entry;
- b. marital status;
- c. highest education level reached;
- d. number of dependents;
- e. various ASVAB subtest scores (in raw score form);
- f. AFQT percentile scores;
- g. groupings based on AFQT scores;
- h. entry pay grade; and
- i. SCREEN score (which is the probability of completing the period of enlistment based on education, AFQT and race).

These variables have all been previously associated with predicting criterion measures in military and non military settings. They are consistent with the three requirements for predictors mentioned in the literature survey.

3. Moderator Variables

As was pointed out earlier, it is suspected that complexity plays a moderating role. This has been accounted for, partially at least, by choosing three ratings which cover a wide span of complexity, and for which different selection procedures will be developed.

Race is another variable which may play a moderating role. Provided there are no adverse impacts for minorities, as measured by hiring rates, it may be possible to apply the same selection procedures across race, to the rating as a whole. Alternatively, selection procedures based specifically on race groupings may produce a better overall result. In any case, previous research emphasises wisdom of taking race into account.

Sex is another variable which is believed to play a moderating role in this research. In this study, there is a sufficient number of white females to perform analyses on two rating groups (these are PN and AT).

4. Screening of the Variables

While the accuracy of data entries into the original three data files is unknown, it is assumed to be very high. However, some cases were identified in the cohort data set which had impossible values. For example, no entrant to the Navy can be less than 17 years old, but some other members were listed with an entry age less than 17. These cases were excluded. Similarly, each of the ASVAB subtests has a known maximum number of items, thus cases with scores higher than these limits were excluded. Because the criterion measures all required that each sailor have the opportunity of serving for at least four years, then those sailors whose enlistment dates precluded serving this period were excluded.

The exclusion of cases for these reasons was 'list-wise'. That is, the case and all its variable values was deleted. The percentage of cases deleted in this way amounted to less than 5%.

5. Estimates of Utility

As was argued in Chapter II, if a significant relationship between predictors and the criterion is found, then estimates of utility for various selection outcomes become important for deciding the optimum cut-off on the predictor to achieve the best overall selection result. This involves determining the cut-off level on the predictor for which total utility is maximised (see Table II and related discussion). This section discusses the derivation of the estimates of utility.

To estimate the utility associated with each of the four possible selection outcomes it is assumed that the Billet Cost Model (see Griffin, 1981, for an explanation of the cost components of this model) provides reasonable estimates. As is discussed by Campbell (1976), it is the intercell ratios which are important rather than the absolute values themselves. The Billet Cost Model provides a method of approximating these ratios.

For correctly predicting a successful sailor, will complete four years of service, the total marginal cost is taken as the best estimate of utility. This is based on the assumption that the Navy compensates sailors at the full value of their marginal product. This is probably a conservative estimate as a recent paper (Butler, 1982) Butler has calculated for the Electronic suggested. Technician Rating, based on the amount of training dollars the Navy spends on these sailors and their expected total service, that the Navy must expect a return in product considerably in excess of rating billet cost. Nevertheless, the total billet cost figure is accepted as the best estimate of utility for this outcome (i.e. correctly predicting success).

An estimation of the disutility of predicting success, when in fact the sailor fails the criterion, is provided by the rating cost estimates of Balis and Clay-Mendez (1982). Basically these data come from the Billet Cost Model, but include only a few of the elements of the criginal model. The Balis and Clay-Mendez cost amounts are the average replacement costs by ratings across so called quality gradings. The cost of the replacement of the highest quality grading is used in the present study as the estimate of the disutility of falsely predicting success.

Relevant product/cost figures are not available for the utilities associated with applicants who are rejected by the prediction device. However, these utilities are derived from the above estimates. In the case of the applicant correctly rejected (he/she would have failed on the criterion measure), it is argued that this selection outcome has the same utility as the disutility associated with the entrant who fails the criterion. Therefore, the 'Balis' cost data is used as the utility estimate for this selection outcome.

The disutility of the rejection of the potentially successful applicant is derived as follows. The disutility associated with this outcome depends on the recruiting market and whether it is easy or difficult to get sailors. In some circumstances it could be that there are so many good quality applicants that rejecting potentially successful applicants has little or no disutility. One extreme value for this utility, therefore, could be zero. At the other extreme, the disutility of rejecting these applicants could be equal in magnitude to the utility of enlisting a successful applicant. The estimate of the disutility of rejecting a potentially successful applicant is taken as the average of these two extremes. The utility values, by rating, are given in Table IV.

TABLE IV
Relative Utilities by Rating

	PREDI *** SUC	CTOR	PREDICTOR *** FAILURE ***		
RATING	Critericn Success	Criterion Failure	Criterion Success	Criterion Failure	
SH	19260	-13077	-9630	13077	
F N	18488	-14495	-9244	14495	
AT	22297	-21210	-11149	21210	

The above utilities are used to derive cut-off scores for all three criteria. For the Length of Service(ICS) criterion they are used directly as the utilities for each of the four possible selection outcomes. On the goodguy and badguy scales they are used as the basis of developing utilities as described below.

TABLE V
Relative Goodgmy Utilities by Rating

	***	PREDICTOR *	** ***	PREDICTOR FAILURE	***
	30 C	riterion 20 1	0 30	riterion 20	10
SĦ	28890	19260 - 13	0 77 - 19260	-9630	13077
E N	27732	18488 - 14	495 -18483	-9244	14495
AT	3 3 4 4 6	22297 - 21	210 -22297	-11149	21210

For the goodguy scale, there are three criterion cutcomes, which for the purposes of establishing a cut-off, gives a total of six possible criterion/predictor combinations. The scale values 20 and 10 are taken to have utilities as shown in the above table. The utility of correctly

predicting a goodguy 30 is computed to be one and a half times the utility of correctly predicting a goodguy 20. As was done in developing the above values, the disutility of rejecting a sailor who would have been a goodguy 30 is taken as the simple average of its extreme values. As can be seen in Table V, the figure for SH is the average of 28890 and 9630.

TABLE VI
Relative Badguy Utilities by Rating

	***	PREDICTO SUCCES:	OR S ***	***	PREDICTO FAILURE	R ***
	30 C	riterio:	n 10	30 C1	riterion 20	10
SH	19260	13077	- 196 16	-9630	-13077	16347
F N	18488	14495	- 21743	-9244	-14495	18119
AI	22297	21210	- 31815	-11149	-21210	26513

Similarly, the badguy utilities, shown in Table VI, were derived starting from those in Table IV. These four utilities represent the outcomes for bad guy criterion scores of 20 and 30. Again, it was figured that the utility of correctly identifying and rejecting badguy 10s was one and a half times the utility of correctly identifying a 20. As for the goodguy scale a simple average gave the utility for the other outcome.

D. SUBJECTS CHOSEN

Within the NPS cohort file there are three possible indicators of a sailor's rating. These are as follows. Firstly, there is a rating recorded at the time of entry. Secondly, when the sailor attempts the rating examination,

the resulting rating is recorded. Thirdly, the DMDC file contains a rating which is the sailor's official designation in terms of rating. These rating variables are known, respectively, as (i) entry rate, (2) examination rate and (3) DMDC rate.

It would be an ideal situation if all three rating variables agreed. However, there is considerable variation in the recordings of rating across these variables. In fact there are a total of seven different combinations when it comes to selecting suitable subjects for a particular rating. These seven possible combinations are shown in the Table VII.

TABLE VII Possibilities for Selecting Subjects within a Rating

,	Rating Indicators					
Category	Entry	Exam	DMDC			
1.	Yes	Y es	Yes			
2.	Yes	ze Y	No			
3.	Yes	Уo	Ύэ́ε			
4.	Yes	No	No			
5.	ИО	СИ	Yes			
6.	Nc	Y es	Yes			
7.	Ис	No	Yes			

Note 1: 'Yes' means that the data file indicated that the sailor was a member of the rating according to the particular rating indicator.

Note 2: In choosing subjects who were representative of rating membership, all the above Categories except # 4. were used as selection criteria.

There are a number of ways for determining which subjects in the listed categories are representative of a rating. One method could be to select only those subjects who were coded in the rating for all three rating selection variables. But this approach would seriously restrict the sample size and ,indeed, it would probably not produce a repesentative sample of the members of the rating.

Of the seven possible categories listed in Table VII, it was decided to initially select on all seven, then to exclude those sailors in category 4. That is, the exclusions were those Navy entrants for whom the only indication of membership of the rating was their entry code. reasoned that while these sailors showed an interest in the rating at some stage prior to entry, there is no indication that he/she had ever had any work experience in the rating. From a concurrent validity point of view, it is necessary for the subject sailers selected for a rating to at least experience it and to have their rating performances influenced by that experience. Presumably the sailors who were coded in the rating through the examination rate or through a DMDC code have had sufficient experience in the rating to be considered representative of the rating.

After these variable screens were applied, and the rating categories identified, subjects were chosen within each rating. These selections are shown by race and sex in Table VIII. As will be seen in the results section, it was decided that the numbers in some of the race/sex combinations were too small to perform realistic analysis. There are sufficient numbers of men in each race grouping. However there are enough white women in only the PN or AT ratings to perform regression and other analyses. Therefore, in all subsequent analyses, three groupings for SH, and four each, for FN and AT, are reported.

TABLE VIII
Rating Selectees by Rating, Race and Sex

RATING	Sex	White	Black	Others	TOTAL
SH	Male	1330	57 <u>2</u>	169	2071
SH	Female	28		2	37
F N	Male	1263	288	112	1663
F N	Female	479	70	15	564
AT	Male	3400	176	114	3690
AT	Female	243	14	3	260

Note 1: The number of personnel are shown in each call.

E. STATISTICAL PROCEDURES

1. General

Now that the criterion, predictor and moderator variables have been indicated and the subjects identified, the statistical procedures will be described. The Statistical Analysis System (SAS) package (SAS, 1979) was used for the analyses, along with some FORTRAN programs written by the author. Sample programs are listed in Appendix A.

2. <u>Descriptive Analyses</u>

These are given by moderator variables, and consist of the ecore and/or mean values on predictor and criterion variables. No statistical tests are applied to these values, since where significant differences occur these are "captured" in the subsequent analyses.

3. Predictive Analyses

For each rating group the following steps are applied:

- a. stepwise regression is used to identify significant predictors;
- b. these significant predictors are used in 'crdinary least squares' regression to develop 'double crossvalidation' validity coefficients; and
- c. regression weights are used to score each case on the respective predictor equations.

When the group size was small, the number of predictors used in the subsequent stepwise analysis, was limited to one variable to every 20 personnel. As a consequence, the 'other' racial category for ATs was limited to the first five variables selected in the stepwise procedure.

4. Estimation of Cut-offs

The 'ordinary least squares' regression equations are used to estimate optimum cut-offs as is described below.

In order to use the utility values to estimate the most ideal out-offs on the predictors, it is first necessary to consider the LOS criterion measure as a dichotomous variable. Thus the sample is divided into two groups: those with less than four years and those with four years or more of service. The two three-point criterion scales were not recoded.

As is illustrated in the Appendix A, during the double cross-validation SAS run on the computer, criterion measures predictor scores and the variable race were output to a separate computer file which could be accessed by the FORTRAN program. The predictor score output at this stage

was derived in the following way: the separate regression coefficients, in each of the cross-validating samples, were used to separately score all the cases producing two predicted scores for each case. A simple average was taken, which represents the predictor score output for later FORTRAN analysis.

The method for estimating the most appropriate cut-off score was explained in Chapter II. In summary, the cutting score chosen is the one which maximises equation 4.1.

 $Uc = Sum \ cf \ (Pr(i) * Ut(i))$ (eqn 4.1)

where:

Uc is the Total Utility for Cutting Score c.

Pr(i) is the Probability of Outcome (i).

Ut(i) is the Utility of Outcome (i).

The results of these analyses will be presented in graphical form. In order to get some frame of reference for evaluating the value of the chosen out-off score, the out-off is expressed as a percentage change from, what is known as base rate. The base rate is simply defined as the value of equation 4.1 when the out-off is set so low that every applicant is accepted. Unfortunately, base rate is not the true utility of the current selection, since it is represents the assessment of the utility only for those who are selected by these procedures, ignoring applicants who are rejected by them. As is discussed in Chapter VI, the overall utility of a selection device must take into account the costs involved in setting up and maintaining the system.

Thus, base rate, as defined here, is a utility estimate which ignores these external costs, but it provides an index figure for judging each potential cutting score. This index for each score is expressed as a percentage change, either positive or negative, from base rate.

V. RESULTS

This chapter presents the thesis results and is organised into three sections. The first is descriptive, providing breakdowns on criterion, predictor and moderator variables. The second section details the predictive results. This includes the variables selected for multiple regression, and also the results of cross validation. In the last section, the results of the application of utilities to estimate appropriate outting scores, are presented.

A. DESCRIPTIVE RESULTS

1. SE Groups

Shown in Table IX are the means and/or frequencies by category for potential predictor and criterion variables across the three male SH racial groupings. It is interesting to note that the length of service is greatest for non-white races. On the other hand, both the goodquy and hadguy scales indicate that proportionally more of the white group have extreme values than do the other two races.

Among the predictor variables there are also some interesting trends. White personnel seem to be younger with lower educational level. However, with some exceptions, the white group's performances on the ASVAB subtests are better. This finding is consistent with this group's better AFQT scores.

2. FN Groups

The same basic trends noted for SH men are evident for FN men, as is shown in Table X. The white men are slightly younger, they exhibit 'good' rather than 'bad'

TABLE IX Criterion and Predictor Values for SH Men

Variable		White	Black	Other
Number of Cases		1330	572	169
Mean LOS: months		44.99	49.31	50.82
Gcodguy Scale	30 20 10	24% 33% 43%	16% 49% 35%	12% 58% 30%
Eadquy Scale	30 20 10	23% 51% 26%	17% 66% 17%	14% 72% 14%
FREDICTOR MEANS Entry Age: years Marital Status Highest Education Dependents (number) AFOI Percentile AFOT Group Entry Paygrade SCREEN Score		18.7 1.3 11.6 .03 46.6 5.3 1.1 80.2	19.4 11.8 10.5 34.7 4.5 79.4	21.0 1.5 12.2 .02 32.3 4.0 1.1 80.2
ASVAB Subtests: Attention to Deration Numerical Operation Autc. Informaticn Attentiveness Scale General Science Electronics Scale Math. Knowledge Space Perception Maintenance Scale Electronic Info. Anithmetic Reasonin General Information Word Knowledge Shor Information Comfat Scale Mech. Comprehension	s	142.74 120.74 109.14 112.9 112.9 112.9 112.9 112.9 112.9 112.9 112.9 112.9	138.53 10.85 10.85 11.52 10.78 11.52 17.19 17.14	12769788590665551449 127697888194965936

Note 1: Goodguy: 30 .. served 4 years, promoted E4 and recommended for reenlistment.

20 .. served 4 years.
10 .. remainder after 20 and 30.

Note 2: Badguy: remainder after 10 and 20. minor negative indicators. major negative indicators.

Note 3: Goodguy and Eadguy percentages sum to 100%.

Note 4: AFOI groups: values 1 to 8, represent categories (in order) 5, 4C, 4B, 4A, 3B, 3A, 2 and 1.

Note 5: Marital Status: married (2); other (1).

TABLE X
Criterion and Predictor Values for PN Groups

Variable		White	MALE Black	Other	FEMALE White
Number of Cases		1263	288	112	479
Mean LOS: months		48.72	49.53	51.94	45.67
Gocdguy Scale	30 20 10	17% 48% 35%	9% 52% 39%	12% 63% 25%	21% 36% 43%
Badguy Scala	30 20 10	17% 72% 11%	9% 81% 10%	14% 79% 8%	225 745 43
PREDICTOR MEANS Entry Age: years Marital Status Highest Education Derendents (number AFCI Percentile AFCI Group Entry Paygrade SCREEN Score	:)	19.8 12.1 12.07 63.0 6.2 1.4 84.3	20.7 12.2 .06 49.3 5.5 82.7	21.1 1.5 12.7 .08 34.8 3.9 1.1 82.3	20.0 11.5 12.52 05 63.9 6.3 1.4 Ncne
ASVAB Subtests: Attention to Detail Numerical Operation Autc. Information Attentiveness Scale General Science Electronics Scale Math. Knowledge Stace Perception Maintenance Info. Arithmetic Reasoni General Information General Information Compation Compation	ng ng	0435697618067680 131112732995033551	140.659 110.659 1119.62820.61 11962820.61 11962820.61	115,697,6897,3964816 12697,6897,3964816	

Notes: As for Table IX.

behaviours, and are better performers on the ASVAB and the AFQT than are the men in the other racial groups.

The scores of the white PN women, not surprisingly, are closest to the scores of the white PN men. The women, however, have considerably shorter average lengths' of service, but exhibit a higher proportion of 'good' rather

TABLE XI
Criterion and Predictor Values for AT Groups

Variable	White	MALE Black	Other	FEMALE White
Number of Cases	3339	172	109	242
Mean LOS: months	53.32	53.39	51.16	45.37
Goodguy Scale 20	183 52% 30%	10% 57% 33%	18% 43% 34%	14% 44% 42%
Badguy Scale 20	19% 76% 5%	10% 83% 7%	16 % 82 % 2%	15% 76% 3%
FREDICTOR MEANS Entry Age: years Marital Status Highest Education Defendents (number) AFOT Percentile AFOT Group Entry Paygrade SCREEN SCORE	19.1 1.3 12.0 72.5 6.7 1.9 86.2	20.5 1.4 12.09 56.5 5.9 1.8 83.6	20.2 1.5 12.08 55.6 1.5 83.9	20.6 1.5 12.5 12.5 79.7 7.0 1.3 Non 9
ASVAE Subtests: Attention to Detail Numerical Operations Autc. Information Attentiveness Scale General Science Electronics Scale Math. Knowledge Space Ferception Maintenance Scale Electronic Information Arithmetic Reasoning General Information Word Knowledge Shor Information Comfat Scale Mech. Comprehension	2574900902362245 15549416142461465	13100215967 1310021133024015967	5513555000816247 53192011110139931411	7504816373766 1400249.6373766 10580696151 111249.63737661 111249.63737661

Notes: As for Table IX.

'bad' behaviours. On ASVAB subtests the women, when compared with white men, are better on some scales (e.g. Numerical Operations and Word Knowledge), poorer on some others (e.g. Auto Information and Electronics Scale) and about the same on the remainder. On educational level, the women have entry educational levels which are within the

extremes of the three male racial categories on this variable. Incidentially, it is the male 'other' race category which has the highest educational level for the PN rating.

3. AT Groups

For the AT rating perhaps the most striking fact is the proportion of ATs who are white (see Table XI, about 92%). This is considerably larger than the equivalent percentages for FN (76% white), and SH (64% white).

The trends in the AT data are consistent with those already mentioned for PNs and SHs. Again, the white group is younger, with lower education level and generally better AFQT scores. However, for AT there are proportionally more Placks in the negative category of the 'bad guy' scale (i.e. those sccres of '10'). Women again have superior performances on some ASVAB scales and, for this rating, they have the highest education level.

B. PREDICTIVE RESULTS

In this section the results of the stepwise regressions, in the form of the variables selected and the signs of the respective coefficients, are presented. These are followed immediately by validity estimates for the predictive models which are constructed. Regression coefficients and statistics are not listed in the body of the thesis: however, interested readers are directed to Appendix B where they are given in full.

1. Length of Service Criterion

For the SH rating, the variables selected as predictors on the length of service criterion by race, according to the sign of the coefficient, are shown on Table XII. Across races all the signs are consistent for predictors

TABLE XII
Stepwise Regression for SH on LOS

V aria tle	White	Black	Other
Entry Age: years Marital Status Highest Education Attention to Detail Auto. Information Attentiveness Scale Ganeral Science Arithmetic Reasoning General Information Mech. Comprehension Combat Scale AFCT Group Entry Paygrade SCREEN Score Dependents (number)	E Se de la Sese de la Caracteria de la C	POOR OR SOR	Ponnannannann Nannoag Nannann Nannann Nannann

Pcs means the coefficient was positive. Neg means the coefficient was negative. In means the variable was not selected.

which entered the equation in more than one racial grouping. Entry age, marital status, educational level and entry paygrade are all positively related to length of service. Longer service is associated with higher values on these variables. Nearly all the predictive ASVAB subtests, and the AFQT groupings, are negatively related to length of service.

As shown in Table XIII, there are some differences in the signs associated with predictors for the PN groups for the length of service criterion. For whites and blacks, higher education is associated with shorter service, while for the 'other' race category, the reverse is true. There are mixed signs for the ASVAB subtests, but the SCREEN score consistently has a positive sign. Some ASVAB predictors for length of service, in the female group, have negative signs. Perhaps this is not surprising, because as was noted earlier, women tended to have shorter service but better ASVAB subtest scores than did men.

TABLE XIII
Stepwise Regression for PN on LOS

Variable	White	MALE Black	Other	FEMALE White
Entry Age: years Marital Status Highest Education Attention to Detail Numerical Operations Electronic Infc. Alithmetic Reasoning Word Knowledge Combat Scale Shop Information AFCT Percentile Entry Paygrade SCREEN Score Dependents (number)	Poes N Poes N Poes Poes N Poes Poes Poes Poes Poes Poes Poes Poes	N N N N N N N N N N N N N N N N N N N	enos Penannanennae N	NN P NN P P NN P NN P NN P NN P NN P P

Note
Pos means the coefficient was positive.
Neg means the coefficient was negative.
In means the variable was not selected.

TABLE XIV
Stepwise Regression for AT on LOS

Variable	White	MALE Black	Other	FEMALE White
Entry Age: years Marital Status Highest Education Numerical Operations Auto. Information Attentivence General Science Electronics Scale Math. Knowledge Space Perception Maintenance Scale Electronic Infor AFCT Fercentile AFCT Group Entry Paygrade SCREEN Score	n segg Neen neegg Neen ncees Neen ncees Nees	n s g Nen nosg n nosg n n n n n g Nen n n	onna pronnacounana o Onna pronnacounana o Onna pronnacounana o	nen nen g Nen nen ne neconnan Nen se

Note
Pos means the coefficient was positive.
Neg means the coefficient was negative.
'n' means the variable was not selected.

Like the SHs, there is a consistency of signs for stepwise entered predictors for the AT rating (see Table XIV) for length of service. However, the direction of the signs, for some variables, is different. For example, educational level for two racial groupings, is now negatively related to length of service. This sign change is carried over into the ASVAB subtests. Being such a technically complex rating, one would expect Science and Mathematics to be positively related to length of service. However both of these scales have negative signs. Consistent with some of the other stepwise regressions, SCREEN score is positively related to length of service.

2. Gcodguy Criterion

Table XV shows the variables selected by race on the goodguy scale for the SH rating. This is the smallest

TABLE XV
Stepwise Regression for SH on Goodguy

V aria ble	White	Black	Other
Marital Status Highest Education Auto. Information Electronics Scale Math. Knowledge Electronic Info. Combat Scale AFCT Percentile AFCT Group SCREEN Score	Pos Pos Neos Neon Rannn	Pos nnnn nn Pos Nes	n neg neos Neos Neos Neos

Pcs means the coefficient was positive. Neg means the coefficient was negative. 'n' means the variable was not selected.

number of variables selected for any of the stepwise regressions. The signs are all consistent with the predictor selections for SH on length of service. However, it is interesting to note that AFQT percentile, for the 'other'

racial category, has a negative sign: lower AFQT percentiles are predictive of better goodquy performances.

TABLE XVI
Stepwise Regression for PN on Goodguy

Variable	White	MALE Black	Other	FEMALE White
Marital Status Highest Education Attention to Detail Numerical Operations Electronics Scale Math. Knowledge Mech. Comprehension Electronic Info. Arithmetic Reasoning Word Knowledge Contat Scale General Information Shor Information AFCT Fercentile SCREEN Score Dependents (number)	POEN S S G G S S S G S S S G G S S G G S S G G S S S G G S	HAND D UND UND DE UND DE UND DE UND DE UND	M MM TO TO THE THOUSE OF THE THE TO TO THE THE THOUSE OF THE	nngg Nnnnnegg Nnnnnnnnnnnnnnnnnnnnnnnnnn

Pcs means the coefficient was positive. Neg means the coefficient was negative. In means the variable was not selected.

As was the case on the length of service criterion for PNs, the level of education is negatively related to the goodguy scale. This is shown in Table XVI. ASVAB subtests selected have varying signs and for two racial groupings for men, the variable 'number of dependents' has entered the regression equation. The SCREEN score for white males is positively related to desirable goodguy behaviors.

For the AT groupings, on the goodgry scale, see Table XVII, there is little consistency regarding which variables are selected. Only three predictors were chosen in more than one grouping, and when this happened the signs are different across groups. It is interesting that for the largest group, white males, not a single ASVAB subtest is selected as a predictor.

TABLE XVII
Stepwise Regression for AT on Goodguy

Variable	White	MALE Black	Other	PEMALE White
Marital Status Highest Education Attention to Detail Attentiveness Scale General Science Electronics Scale Maintenance Scale Flectronic Info. Word Knowledge AFCT Percentile AFCT Group Entry Paygrade SCREEN Score	Pos Fos n n n n n n n n	nense Nense Nenennoen Nenennoen Nenennoen	EROPERROS NPN P. NPN P.	n n n n n e o s n n n n e o s n n n n n n e o s n n n n n n n n n n n n n n n n n n

Pos means the coefficient was positive. Neg means the coefficient was negative. In means the variable was not selected.

3. <u>Padguy Criterion</u>

The SH rating, see Table XVIII, on this scale, shows a consistent trend for education to be positively related

TABLE XVIII
Stepwise Regression for SH on Badguy

Variable	White	Black	Other
Entry Age: years Marital Status Highest Education Numerical Operations Autc. Information Electronics Scale Arithmetic Reasoning Combat Scale Word Knowledge AFCT Percentile AFCT Group Entry Paygrade SCREEN Score	SSS Q C SSS	N C NP THE THE	nnonen nonen n Ponen sonen n Ponen n

Pcs means the coefficient was positive. Neg means the coefficient was negative. In means the variable was not selected.

for all three racial groupings. Since the goodquy and the badguy scales are both scored in the same direction (i.e. the 'good' and 'bad' ends have the same numerical values), this means that higher education is associated with 'non-bad' behaviours. Several ASVAB subtests were selected, with differing signs. Entry paygrade and SCREEN score are both positive.

TABLE XIX
Stepwise Regression for PN on Badguy

Variable	White	MALE Black	Other	FEMALE White
Entry Age: years Highest Education Attention to Detail Auto. Information Space Perception Electronic Info. Anithmetic Reasoning Shor Information AFCI Fercentile AFCI Group Entry Paygrade SCREEN Score Derendents (number)	nnnnnonnnenoo Pranenoo Neo PP	Nencenneen s Ponneen s Ponneen s	NA N	N n n n n n n n n n n n n n n n n n n n

Pcs means the coefficient was positive. Neg means the coefficient was negative. 'n' means the variable was not selected.

For FNs, see Table XIX, entry age entered the regression only for blacks, and it has a negative coefficient. Elucation level is selected for two groups, the 'other' male race category and white women, and for both, the sign is positive: higher education level implies fewer undesirable behaviours. There are a variety of ASVAB subtests selected with differing signs, and it is interesting to note that for three of the four groups the number of dependents entered the regression. For white males, having more dependents suggest fewer negative behaviours, while for the 'other' races and white females, having more dependents is associated with more negative behaviours.

TABLE XX
Stepwise Regression for AT on Badguy

Variable	White	MALE Black	Other	FEMALE White
Entry Age: years Marital Status Highest Education Auto. Information Electronics Edge Math. Knowledge Electronic Info. Arithmetic Reasoning Comfat Scale AFOT Group Entry Faygrade SCREEN Score Dependents (number)	POPS POPS POPS POPS POPS POPS POPS POPS	nn no ne no sog ne no ne no ne no ne no ne no ne	HERECON SECULATION OF SEC	Pos neg neg nos neg neg neg

Pos means the coefficient was positive.
Neg means the coefficient was negative.
'n' means the variable was not selected.

The final stepwise table. Table XX, shows the selected predictors for AT groupings on the badguy scale. The most consistent subtest finding across groupings is the entry paygrade, which has a negative sign. This variable is selected in three groupings and implies that the higher the entry pay grade the more likely are negative type behaviours.

4. Validity Estimates

Shown in the next series of tables are double cross-validation coefficients by rating, grouping and randomly selected sample. They are based on forming predictor equations from the stepwise procedures for each individual group. Validaties for the "men all" category were the result of forming regression equations on all predictors which had been independently selected by male race groupings. The average validaties given were calculated using Fisher's transformations for determining average weighted correlation coefficients (see McNemar, 1963, pp. 139-140).

a. Length of Service

The length of service validities by individual sample are shown in Table XXI. All of them are significant except for the 'other' male race grouping, in both the PN and AT ratings, and the women in the AT rating. Most of the validities are greater than .20, but only one (AT blacks in Sample 1) is in excess of .40.

TABLE XXI

Length of Service Validities by Rating and Sample

!	l	Sample	1	<u> </u>	Sample	2 1
RATING	ກ	Γ	p <	n	<u> </u>	þ <
SH MEN All White Black Cther	1016 671 262 91	.2748 .1893 .3375 .3053	.0001 .0001 .0033	10 22 635 296 78	.2806 .2330 .2581 .3640	.0001 .0001 .0001
FN MEN All White Black Cther WOMEN White	817 606 158 66 241	.2731 .2881 .3519 .0546 .2060	.0001 .0001 .0001 .6634	783 620 125 46 238	.2785 .2894 .2963 .1294 .2939	.0001 .0001 .0008 .3919 .)001
AT MEN All White Slack Other WCMEN White	1824 1676 91 58 124	.2530 .2617 .4227 .2393 .1219	.0001 .0001 .0001 .0705 .1771	1799 1678 85 85 116	.2803 .2656 .3338 .2125 .2399	.0001 .0001 .0019 .1304 .0095

t. Goodguy Criterion

For each rating as shown in Table XXII, some of the validities by groups are statistically insignificant. The overall magnitude of these validities, compared to those for the length of service criterion, are smaller. The maximum validity is .378, while a number of them are less then .10.

TABLE XXII

Gcodguy Validities by Rating and Sample

	!	Sample	1	<u> </u>	Sample	2
RATING	n	ŗ	p <	n	ľ	p <
SH MEN All White Elack Cther	10 16 681 262 83	.1052 .1586 .1268 .3673	.0008 .0001 .0403 .0006	1004 649 296 73	.1373 .2176 .1593 .3916	.0001 .0001 .0060
FN MEN All White Elack Other WCMEN White	8 17 606 160 56 241	.0914 .1695 .1579 .0987 .1752	.0090 .0001 .0462 .4306	783 520 128 46 238	.0666 .1464 .1179 .2631 .1799	.0627 .0003 .1852 .0773
AT MEN All White Black Cther WOMEN White	1824 1712 91 58 125	.0421 .0895 .3046 .2577 .0004	.0723 .0002 .0033 .0508 .9969	1799 1688 85 52 118	.0339 .0303 .3503 .3464 .0514	.1503 .2128 .0010 .0119 .5807

TABLE XXIII

Eadquy Validities by Rating and Sample

	!	Sample	1	l	Sample	2
RATING	n	Ε	p <	n	r	p <
SH MEN All White Elack Other	10 16 571 270 91	.1713 .1967 .1178 .3652	.0001 .0001 .0531 .0004	1004 635 302 78	.2062 .2589 .1890 .3254	.0001 .0010 .0037
PN MEN All White Elack Other WOMEN White	817 606 160 66 241	.0963 .1452 .2852 .1102 .1815	.0059 .0003 .0003 .3784 .0047	783 620 128 46 238	.0980 .1174 .2049 .1717 .1896	.0061 .0034 .0204 .2538 .0033
AT MEN All White Elack Cther WOMEN White	1824 1676 91 58 125	.2293 .2513 .2741 .3728 .2688	.0001 .0001 .0086 .0040	1799 1664 85 52 118	.1999 .2086 .3670 .4464 .2400	.0001 .0001 .0006 .0009

c. Badguy Criterion

For this criterion, validities seem to be higher than for the good guy criterion, as shown in Table XXIII. Only four of the 28 sample validities are insignificant at the .01 level of significance. The highest validity of any presented so far (.4464) occurs for one of the 'other' race samples (for ATs).

5. Average Validities

Using the Fisher method, average validaties were calculated from the validaties just presented and these are shown in the next two tables. Table XXIV, shows the average

TABLE XXIV Average Validities by Group and Criterion

CRITTERION

		CRITERION	
RATING	LCS	Goodguy	Badguy
MEN All White Elack Cther	. 2777 . 3107 . 2959 . 3327	.1212 .1876 .1441 .3787	.1879 .2272 .1556 .3470
FN MEN All White Black Cther WOMEN White	. 2757 . 2888 . 3277 . 0852 . 2502	.0793 .1578 .1412 .1669 .1775	.0971 .1312 .2500 .1354 .1855
AT MEN All White Elack Other Women White	• 2666 • 2737 • 3801 • 2262 • 1796	.1380 .9602 .3268 .3002	.2148 .2301 .3197 .4082

validity across each of the samples by racial groupings and ratings. The largest average validity occurs for the AT rating (.41) in the 'other' racial group on the bad guy criterion. The second highest validity is also for an AT rating (.38); it is for the black racial group.

Table XXV, lists grand average validaties for men, for rating by criterion. For length of service, all three rating average validaties are in excess of .25, with a slightly increasing trend in size of coefficient as the

TABLE XXV
Validities for Males by Rating and Criterion

	CRITERION							
	RATING	ICS	Goodguy	Badguy	ĺ			
	SH	. 2515	.2046	.2277				
	FN	. 2707	. 1562	. 1517	ĺ			
i	AT	. 2762	.0867	.2450				

rating complexity increases. For the goodguy scale, the rank order of validities by magnitude is reversed: validities decrease with rating complexity. The average validity for ATs on this criterion is particularly small. There is no consistent ranking by complexity for the other criterion, badguy. The PNs validity remains about the same as for the goodguy scale, as does the SH validity. The average validity for the ATs on this scale has increased considerably, up to .245.

C. ESTIBATING CUT-OFFS

The method for estimating the cut-offs was described in the previous chapter. However, to illustrate the method, the calculation of the first few values in Figure 5.1 is given here. It should be noted that most of the figures in this section have two or more graphs drawn on them. This is done in the interests of economy rather than to imply they are related: each graph is based on a separate predictor criterion relationship.

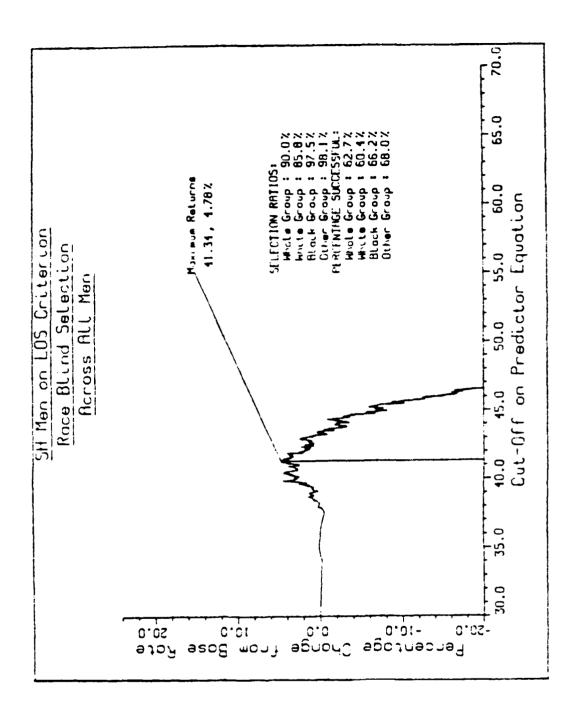


Figure 5.1 Race Blind Selection for SB on LOS.

The figures give several pieces of information, the most important of which is the line repesenting percentage change from the base rate, for various potential predictor cutting scores. For Figure 5.1, the base rate is determined as follows. Of the 2,020 male SH personnel 1,235 served four years or more, while 785 served for less than four years. Converting these values to probabilities gives a chance for a male SH meeting the criterion of completing four years or more years of service, and 38.86% chance of failing it. Using the utility values given in Table IV and equation 4.1 we now calculate the base rate as the sum of the products of the chances of an outcome by its utility. The SH length of service base rate is 6,693.39. reader will recall, this is the value for selection set so low that all personnel are accepted. The first potential cutting score is the smallest predictor score from the sample of 2020. This score is 34.1208 and occurs for an SH who passed the four year criterion. With the cutting score set at this value, there are 1,234 sailors out of 2,020 who are successful and who would be correctly predicted as successful based on this predictive relationship. sailor cut of 2,020 would have been predicted to have failed the criterion for this cut-off, while 785 would have been incorrectly predicted to pass. The numbers are converted to probabilities to give 61.09%, .05% and 38.86%, respectively. These three probabilities have utility values of 19,260, -9360 and -13077. Summing across the products of probabiligives a total value ties and utilities of 6679.23. Therefore, the return over base rate for a cut-off at 34.1208 is a -0.21% change. There is a slightly negative return over base rate for this cutting score. On Figure 5.1 the values 34.01208 and -0.21 are plotted as the first cut-cff coint and its respective return. The next highest predictor score, 34.4115, is choosen. This happens to be for

a sailer who fails to meet the criterion. At a cutting score of 34.4115, the probabilities and their respective utilities are: 61.09% and 19,260; .05% and -9360; .05% and 13077; and, 38.81% and -13077. Summing the product of these values gives a total utility of 6685.70, which is a -.11% decline ever base rate. Therefore the next point plotted is 34.4155 and -.11. This process is repeated for all 2020 predictor scores keeping track of the value of the maximum return over base rate and its respective cutting score. Note in Figure 5.1 that values of base rate return which are less than -20.0% are not plotted.

Also shown in the figures are two other pieces of information. The selection ratio is defined as the percentage/ proportion of personnel selected out of the total who apply. The values shown in the figures are selection ratios at the cutting score which maximise total return over base rate. On each figure selection ratio is shown either by race (for or rating (for women). The other statistic is the critericn proportions for those who are predicted to be successful. For Length of Service, the percentages who serve for at least four years are given, while for the other criteria the percentages by criterion categories are shown. This latter statistic gives an idea as to the expected performance of the group selected for this cutting score, while the selection ratio indicates how the selection device will impact the applicant group at this score.

1. ICS Criterion

As can be seen from Figure 5.1, the optimum sut-off score for the SH rating produces almost a 5% increase over the base rate utility. The predictor equation here is for male SHs based on all those variables selected for men in Table XII. The Figure shows the selection ratios and success percentages for the optimum cut-off score by race.

At this cut-off, white SHs have the smallest selection ratio of any race, but they also have the smallest percentage of success for those who would be selected. The majority of the 'black' and 'other' racial groups are selected and the highest percentages are successful in these groups.

Shown in Figure 5.2 are the cutting scores derived for race specific predictor equations. The best expected return is for the 'other' group (28.8%), followed by whites (8.9%), and blacks (1.7%). Set at these respective cut-off levels, 85% of the 'other' group are selected with about 76% of them serving beyond four years. For whites, 95% would be selected, with 59% successful. For blacks, virtually all (99.5%) would be selected, with a 67% success rate.

In comparing these figures for SH on LOS (Figures 5.1 and 5.2), several trends are apparent. The 'race-blind' selection seems to be most severe, in terms of selection ratio, on the white group. When the white specific predictor equation is employed, more whites are selected, for a slight decline (about 1.5%) in the percentage success. For the black group, a race specific predictor also improves the selection ratio, with a slight decline in the success rate. However for the 'other' race group, the race specific predictor equation markedly reduces the selection rate (down by about 13%), but boosts the success percentage by almost the same amount.

Figures 5.3 and 5.4 show LOS results for the PN rating. Race blind selection produces about a 6% improvement over base rate. As for SHs, the race blind predictor is most severe on the white group in terms of selection ratio. However, for those whites selected, their success percentage is better than the black PNs. While the blacks are the least successful for this predictor equation, about 97% of them are selected. At the out-off level for this race blind equation, 100% of the 'other' racial group is selected, with about 76% of them being successful.

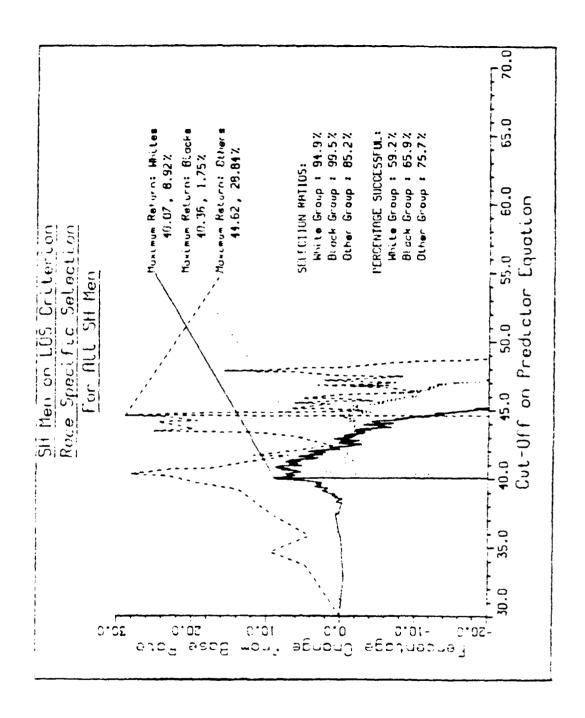


Figure 5.2 Race Specific Selection for SH on LOS.

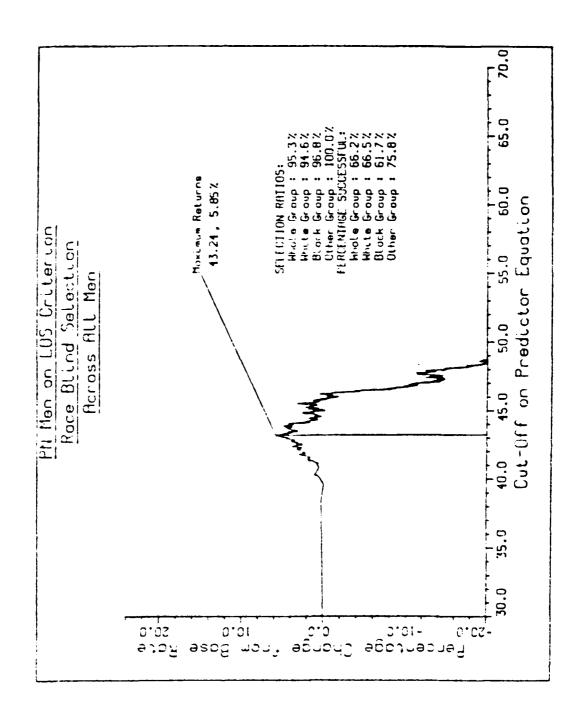


Figure 5.3 Race Blind Selection for PN on LOS.

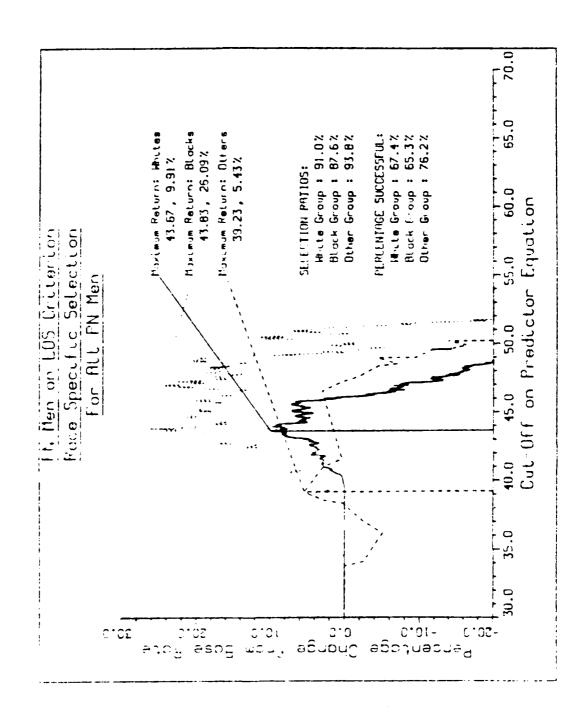


Figure 5.4 Race Specific Selection for PN on LOS.

The return for individual race selection equations (Figure 5.4) is greater than from the race blind selection for each race. Blacks show the biggest improvement, with a reduction in their selection rate, down from 97%, but an improvement in their success percentage. Next come whites whose selection ratio has also been reduced but their success percentage improved. This is followed by the 'other' racial group, who are now selected at a rate of about 94%, with a slight increase in overall success percentage. Therefore, for all racial groups, race specific predictors have lead to a decline in selection rate, but improvements in overall utility and success percentages.

As can be seen in Figures 5.5 and 5.6, trends are not true for ATs on LOS. Race blind selection, lead to about a 3% improvement on base rate, with all races being fairly close together on selection ratios and success percentages. The white race is marginally ahead on both of these indices. In Figure 5.6, about the same values apply for whites on race specific selection. However, specific equations for blacks and for the 'other' racial group have determined optimum cutting scores for which all of the cases would be rejected. That is, no one is selected and thus no one is successful. This result may have been expected for AT 'others' because of the insignificant validity coefficient for this group (see Table XXI). validity suggests that any optimum cut-off value is likely to be spurious rather than meaningful. However, shows a highly significant validity for blacks. deriving an optimum cut-off which rejects everyone is cartainly not expected.

Shown in Figure 5.7 are the LOS results for white women FNs and ATs. Each optimum cut-offs have improved the base rate figure (24% for PN and 36% for AT). While the selection ratios are lower than their male counterparts, so

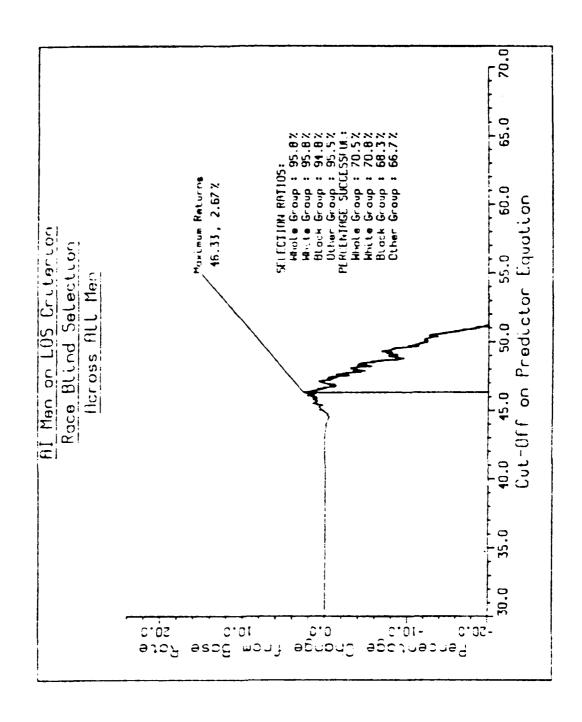


Figure 5.5 Race Blind Selection for AT on LOS.

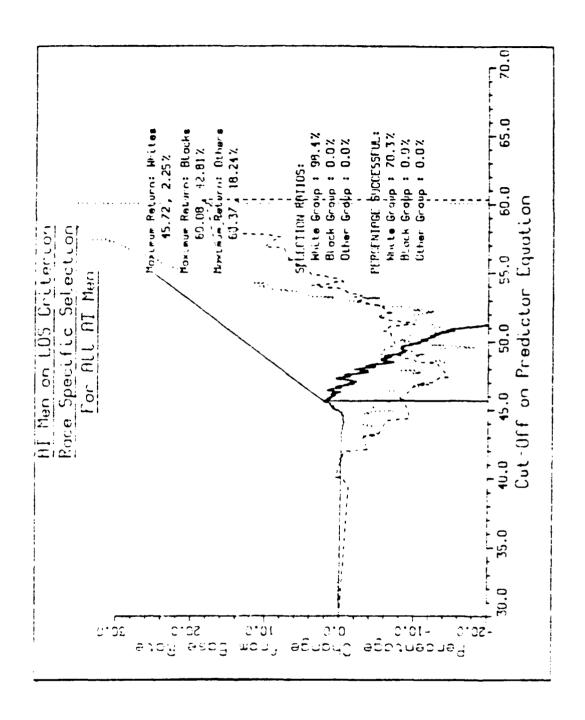
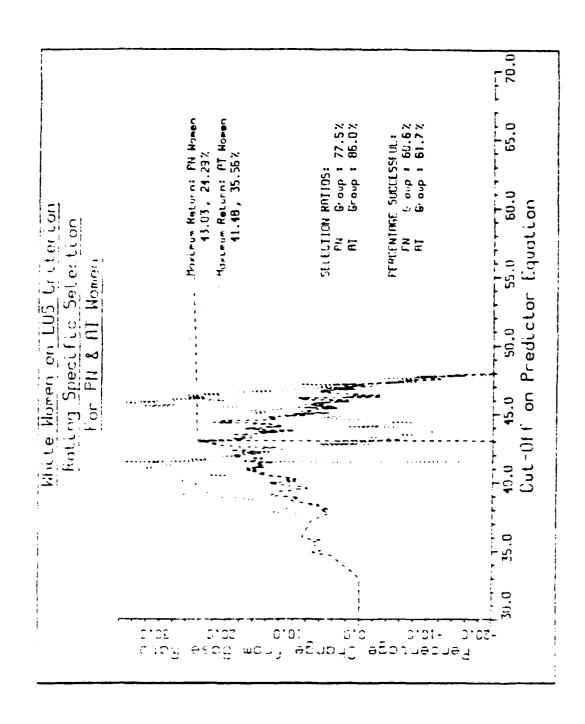


Figure 5.6 Race Specific Selection for AT on LOS.



Pigure 5.7 Rating Selection for PN & AT Women on LOS.

too are their success percentages. This is not so surprising, since Tables X and XI show women have the smallest average LOS of all groups in these ratings.

2. Gccdguy Criterion

Shown in Figures 5.8 to 5.14, are the results of the analyses on this criterion.

For the SH rating, race blind selection leads to a selection ratio for whites that is the smallest of all three races. However, as can be seen from Figure 5.8, the optimum cut-off yeilds a return of just over 2%. For race specific selection (see Figure 5.9), the returns for whites and 'other' races are well over 30%, but in this case, only 65% of the white group would be selected. The selection ratios are reduced for all races, compared with race blind selection.

The returns for the PN rating on the goodguy scale, follow a similar pattern as can be seen from Figures 5.10 and 5.11. However, for this rating, blacks in the race blind selection, have the smallest selection ratio, but they also show the largest return (42%). Indeed, race specific selection has improved their position in selection ratio terms from the lowest for race blind selection to the midpoint on race specific selection.

for ATs on the goodguy scale, race blind selection does not produce any positive return over base rate (see Figure 5.12). The return for race specific selection (Figure 5.13), is substantial for blacks and 'others' (38% and 76%, respectively) but a little more than 1% for whites. For the 'other' race group, this high return is achieved with a selection rate of about 61%. For female PN and AT on the goodguy criterion, see Figure 5.14, the returns are almost 600%. The cut-off for AT seems unrealistic, since only 14% are selected, whereas for PN the selection rate is 75% means that a reasonable proportion will be selected.

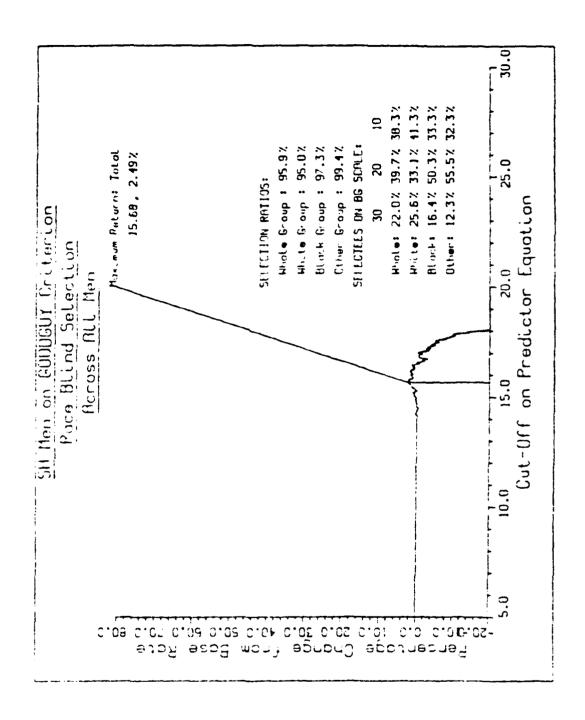
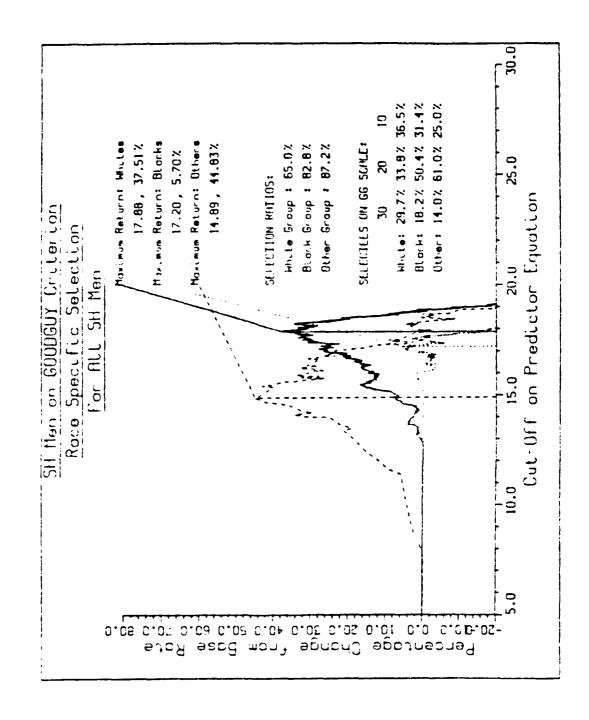


Figure 5.8 Race Blind Selection for SB on Goodguy.



Pigure 5.9 Race Specific Selection for SH on Goodguy.

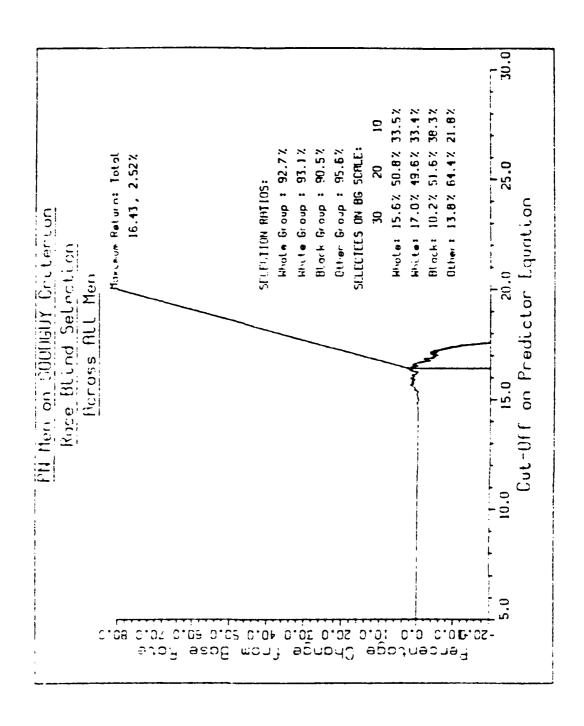
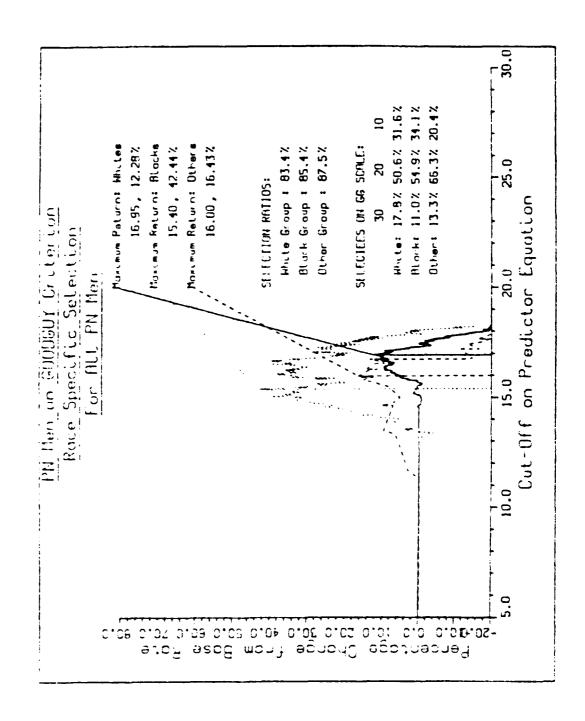


Figure 5.10 Bace Blind Selection for PN on Goodguy.



Pigure 5.11 Race Specific Selection for PN on Goodguy.

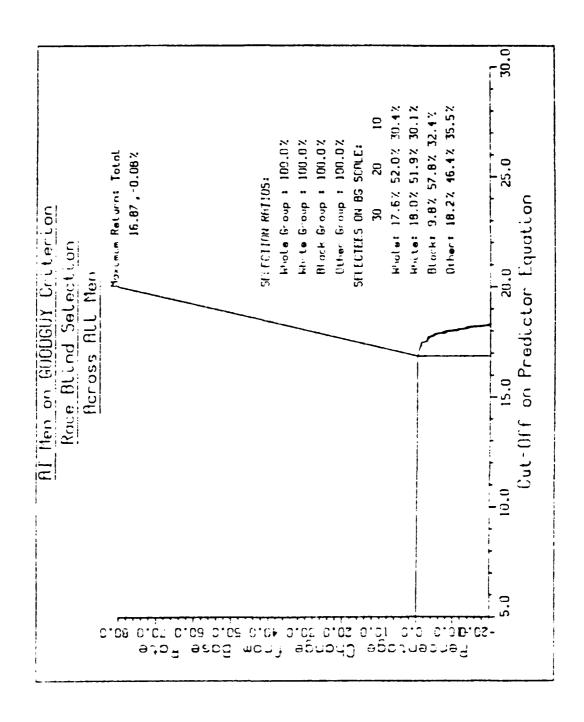


Figure 5.12 Bace Blind Selection for AT on Goodguy.

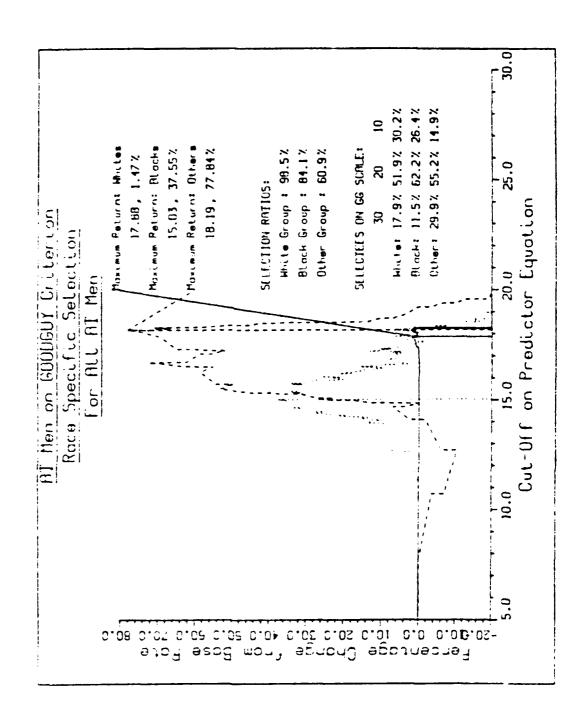


Figure 5.13 Race Specific Selection for AT on Goodguy.

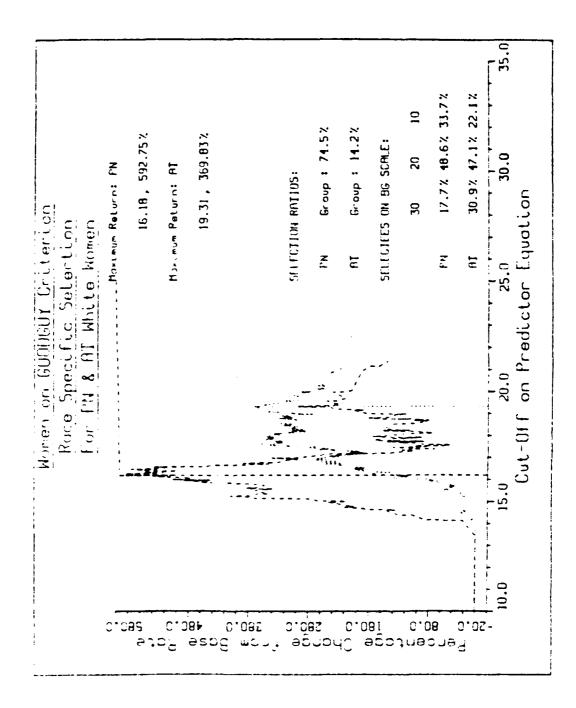


Figure 5.14 Rating Selection for PN & AT Women on Goodguy.

3. <u>Eadquy Criterion</u>

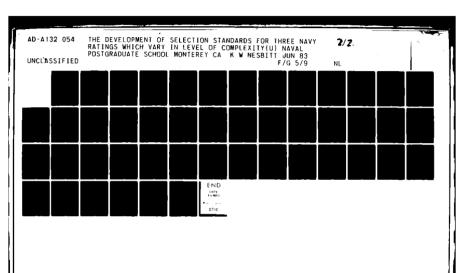
The seven Figures 5.15 to 5.21, show the results for this criterion.

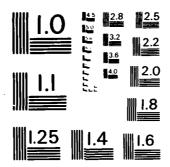
Race blind selection for male SHs, see Figure 5.15, shows about a 3% improvement over base rate. White men are the race which have the lowest selection ratio in race blind selection. Moderate improvement for whites and the 'other' racial group occur for race specific prediction (Figure 5.16).

For PN, see Figures 5.17 and 5.18, there is virtually no effective return using the badguy criterion. While the 'other' racial group shows only about a 6% improvement over base rate, the selection ratio of 98.2% is extremely high.

As can be seen in Figures 5.19 and 5.20, for AT men no selection device has been found that produces an improvement over the base rate.

For women on this criterion, Figure 5.21 shows a positive (AT) and a negative (PN) return. The AT return is difficult to interprete and results in a 99% selection ratio.





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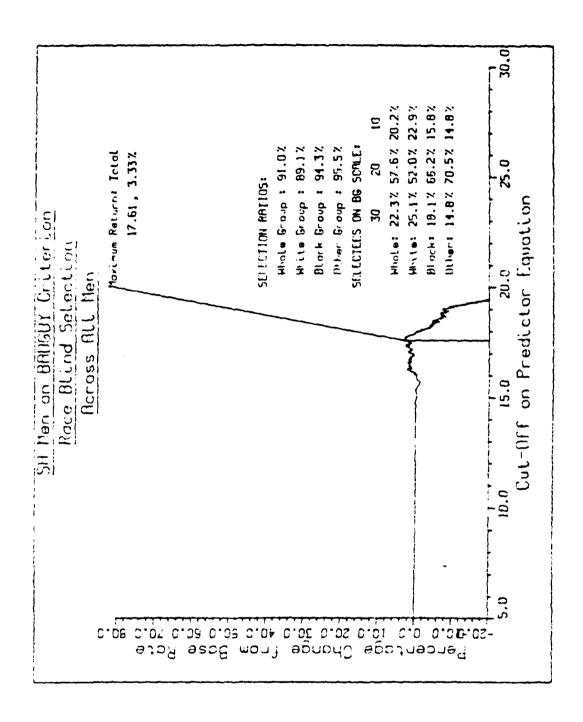
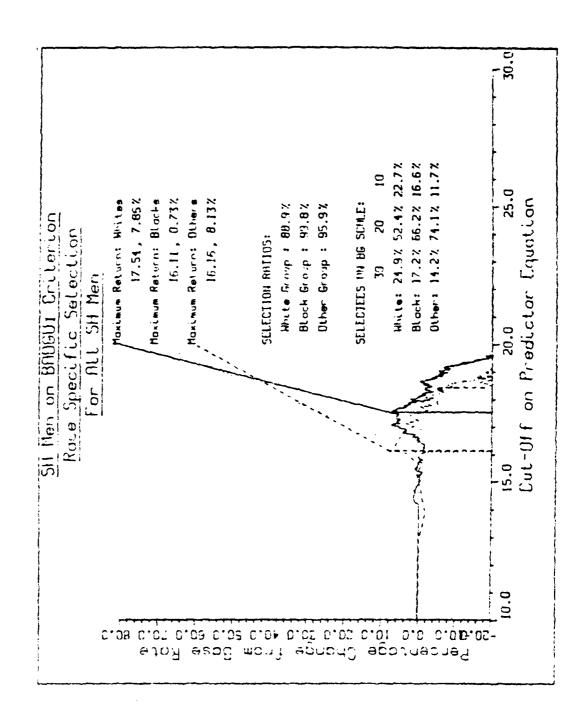


Figure 5.15 Race Blind Selection for SH on Badguy.



Pigure 5.16 Race Specific Selection for SH on Badguy.

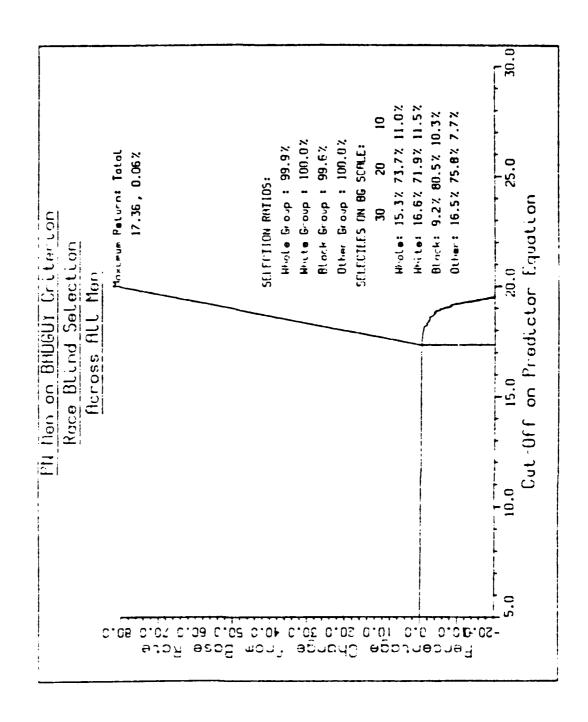
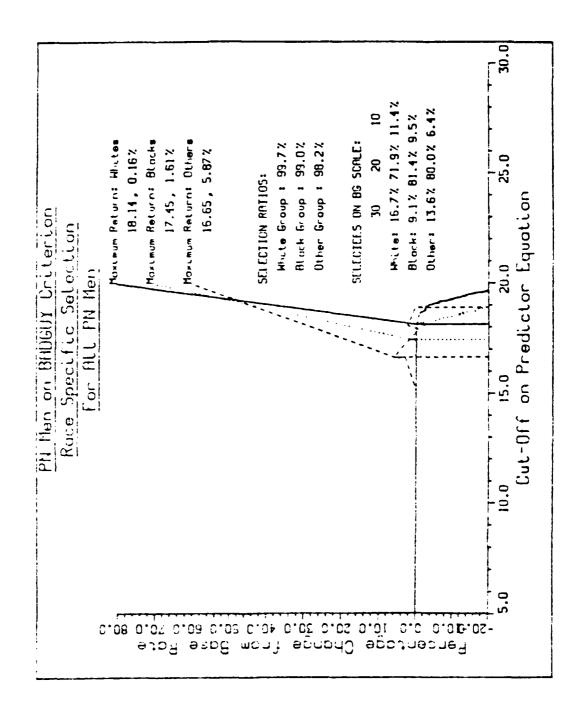


Figure 5.17 Race Blind Selection for FN on Badguy.



Pigure 5.18 Race Specific Selection for PM on Badguy.

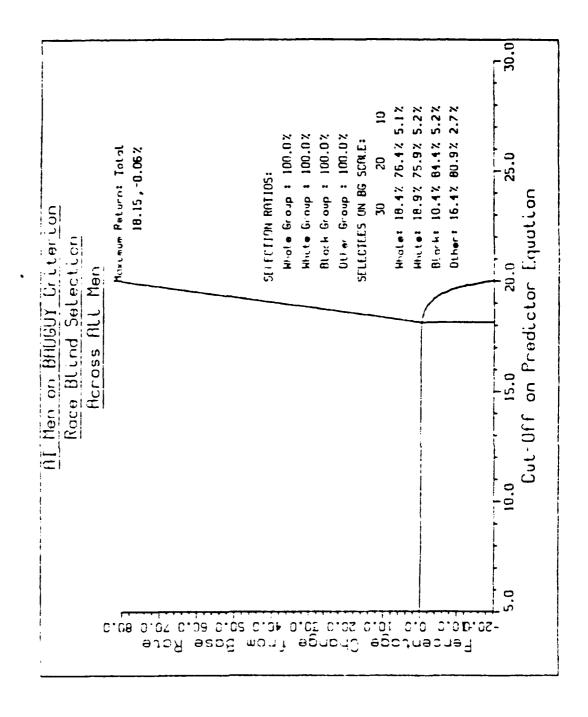


Figure 5.19 Race Blind Selection for AT on Badguy.

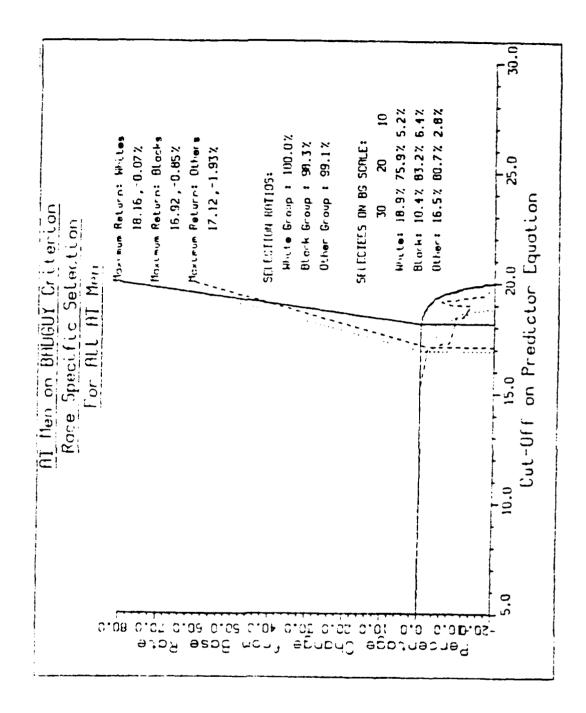


Figure 5.20 Race Specific Selection for AT on Badguy.

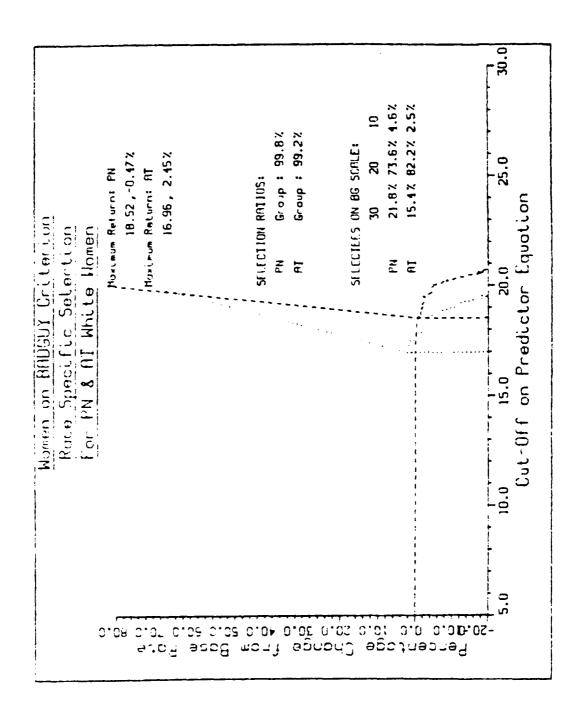


Figure 5.21 Rating Selection for PN & AT Women on Badguy.

VI. DISCUSSION

This chapter is divided into two sections. The first deals with the results in relation to the research hypotheses. The second discusses other findings.

A. RESEARCH HYPOTHESES

The stepwise regressions, particularly on the LOS criterion, tend to support the first hypothesis which predicted that entry age, education level and ability tests would be significant predictors of performance. Entry age was not selected for every regression, but it proved to be significant for at least one of the race/sex groupings on LOS for each rating. The relationship between age and the criterion was always positive: longer service is indicated for greaterage at entry. Educational level was often selected toc. However the direction of the relationship was not as consistent. For some groups it was positive and for others negative. There seems to be a trend for the coefficient for education level to be positive in low complexity ratings and negative in the higher complexity ratings. This finding could be a related to restriction of range problems which were noted in the literature review. At least some of the ASVAE sores entered every stepwise regression. signs were not consistently positive or negative.

At least for one criterion (i.e. LOS) the validity coefficients when averaged across groups, show a trend to increase with jcb complexity, a relationship predicted by the second hypothesis. On the whole, the validity coefficients are large enough to suggest that the predictor equations are sufficiently powerful to improve selection on

these three criteria. Comparing criteria, LOS has the highest validity.

P. OTHER FINDINGS

The descriptive results yeilded several general findings. They seem to support the view that whites have surerior performances on psychometric tests of ability. On each on the ASVAB subtests, with one or two exceptions, the white group in each rating has performed better, in terms of raw score, means than the other two race groups. On the other hand, this race, for each rating, is younger at entry and (perhaps therefore) has the lowest educational level, with the smallest proportion being married variables which might he expected to be associated with lower ASVAB scores. the SH and PN ratings, the white group has a shorter LCS, but with the highest proportions of in-service 'good' and 'bad' behaviours. For the AT rating, blacks and whites have about the same LOS (blacks serve slightly longer), whites are more likely to have positive goodguy perfor-The white women have similar performances as white men on both the predictor and criterion variables, except that they have the shortest LOS and generally exhibit the smallest percentage of 'bad' behaviours.

The estimation of cut-offs, for all criteria, has supported the notion that the significant relationship between predictors and criteria can be useful in a selection setting. As measured against the concept of base rate, many groupings within ratings can be better selected.

The findings of this reseach support the current US Navy selection procedures. For every stepwise regression in which it was selected, the SCREEN score (which is derived directly from the current procedures) was always positively related to the criterion.

From the averaged validaties it appeared that the LOS predictors provide a means of increasing the selection utility across most trade, race and sex groupings. However, when optimum cut-offs were derived it was apparent that virtually none of the AT relationships would yield reasonable utility improvement over the base rate. For the other ratings it appears that, except for SH blacks, there are positive returns to be obtained through selection based on the identified predictor/criterion relationships. highest returns are for PN and AT women. This is despite the fact that some predictor equations had validities which were small, and in some cases statistically insignificant (e.g. FN 'cther' and AT women). Apparently, although the validity coefficient was insignificant the relationship was still strong enough to yeild a cutting score which improved overall utility.

The other two criteria, although they did not have dcuble cross-validation coefficients which were as large as for LCS across groupings of personnel, also apparently are predictable from pre-entry information. However, when the cut-offs were determined, the goodquy criterion appeared to have potential for improving selection in all groupings, with the exception of AT white women. For only three groupings on the badguy criterion was a cut-off found which returned better than 5% over base rate. In a real sense these seem more useful criteria than LOS, since they appear to capture more relevant information relating product and/or cost. The extent that personnel are promoted during a period of service and the extent to which costly negative behaviours can be avoid are significant contributers to the orgamisations overall effectiveness. It would appear that the decision centered validity approach has revealed a number of potentially useful relationships which may be used for better selection.

The race blind selection cut-offs provided some interesting findings. For the SH rating in particular it seems clear that the white race would have greater proportions rejected than the other two racial groups if the race-blind prediction device was used. However, as has already been mentioned, more of the white SH personnel should be rejected since they exhibit poorer criterion performances than the 'other' races. For the PN rating, race blind selection also selects the smallest percentage of whites. This is a more curious result since for both ratings, most of the personnel were white and one would expect that they would dominate the regressions to the extent that greater proportions of whi would be selected than would the other races. specific selection the white selection ratio at the opti cut-off is no longer the lowest for the LOS, but it is the goodguy scale. It appears that whites do not perform is well as the other races in these ratings (PN and SH).

The decision centered validity approach has turned out to be a powerful tool for the present research. allowed a means of directly demonstrating the usefulness of a composite predictor in a future selection role. appropriate utility values can be derived, it is obvious that the methodology employed in this study could be broadened. For example, in evaluating which cut-off to apply, the criterion can be segmented into as many separate categories as is required to reflect the different individual raturns for the pairs of predictor outcomes. At one extreme in criterion categorisation is the LOS methodology employed here, in which the criterion was considered as a dichotomy. At the other is the situation in which each and every subject in the research could be individually evaluated in utility terms. For example, for every subject a value could be placed on the service he/she provided, and an estimation could be made as to the value to the Navy (positive or

negative) in the case that this person was not enlisted because of selection screening. These two values, which would probably not be the same, could be taken as predictor accept and predictor reject utilities and then used in exactly the same manner as in this study to determine an optimum predictor cutting-score which maximises total utility for all subjects. If the utility values were accurately derived from costs and products, then it may be possible to express selection utility in dollars per enlistee, which would be a readily acceptable means of justifying a selection procedure.

In applying the regression equations developed here to an actual setting, there are two other important considerations. These are both related to decisions that the Navy would make with respect to implementation. In this thesis, there was no external Navy decision as to which criteria of performance is the most useful for personnel selection. Length of service is widely used as a measure of good/ desirable behaviour, but for some employments, it may not be so important. Indeed for some employments, "short" tenure may even be a goal (e.g. those employment, which serve primarily to prepare personnel for later jobs). any attempt is made to implement these results the criterion would need to be specifically defined. It could be that there are other more pertinent criteria than those employed in this thesis.

The other issue about implementation is related to costs and has two aspects. First, the present study estimated costs and utilities through published cost data. These figures were clearly and directly relevant for the LOS criterion, but even in this application assumptions were made as to the relative value of the predictor/criterion outcomes. There are, perhaps, better accounting or economic procedures to estimate these values than was used here. An

evaluation of the utilities employed here is needed before implementation of these selection devices. Second, the other aspect of cost is directly related to implementation. determining the optimum cutting-score on the predictive device for this study and expressing this as a percentage return over base rate, it is clear that it does not take into account any costs associated with implementation. A cost-benefit type of analysis would be necessary to weight up the total costs of implementing and maintaining the procedures, against the expected return in terms profitable selection. The data presented here are relevant to this type of analysis, but returns over base rate, example, would need to directly take account of the numbers upon which each of the graphs in the previous chapter was based on. Some of the very high returns over base rate may be seen in a different light when the numbers of cases is taken into account. Ease rate is a utility per enlistee: if the applicant pool is small, total utility (the product of the number of enlistees by utility per enlistee) might not be so high. If the cost of implementing and maintaining a procedure was very high, then the expected return over base rate would have to be even higher, so as to produce a positive return.

VII. SUMMARY AND CONCLUSION

This thesis set out to develop selection standards for three US Navy ratings which varied in terms of their The selection literature was complexity requirements. extensively reviewed. This gave direction as to the variables likely to be predictive of subsequent performance as well as some quidlines for choosing appropriate criteria to predict. The method of evaluating the predictors in forecasting subsequent performances took two forms. The first was the traditional method for estimating validity coefficients through double cross-validation. The second applied utility estimates to various predictor/critetion outcomes so as to determine the cut-off score on the composite predictor which maximised overall utility. These analyses were done controlling for the effects of such moderator variables as job complexity, race and sex.

Within each rating a number of useful predictor/criterion relationships have been found. These relationships were shown to be valid and cutting scores derived which maximised total utility of the selection device. Some results suggest a significant race component in performance. This finding suggests the white race does not perform as well as the other races in the SH and PN rating. The signs on some ASVAB coefficients indicate negative relationships with criterion scores. This result was not expected.

This research has illustrated the relationships that exists in the data and the use to which they could be put for selection purposes. Implementation of any of these relationships as a selection tool requires a confirmation that the criteria used here are relevant and useful to personnel managers in the US Navy. In addition, the current

analysis has been performed without any consideration of the possible costs of the introduction and maintenance of the selection procedure. It would be necessary to conduct a cost benefit analysis to determine if the total expected returns from implementation were sufficient to compensate for the costs involved and therefore to produce a positive total return. To do this properly requires Navy decisions about the appropriateness of the criteria to be predicted and the utility and disutility values to be assigned to correct and incorrect predictor/criteria classifications. Such a cost benefit analysis has not been attempted here, but is relevant for future research.

<u>APPENDIX A</u> EXAMPLE PROGRAM LISTINGS BASED ON PROCESSING SH

This appen ix gives sample listings for the programs used to analysise the SH rating. These exactly parallel analyses for the other ratings.

1. SAS Program to Select SH from NPS Cohort Data

Cn the following pages (Table XXVI) is a listing of the SAS jcb to select out the SH from the Cohort File of 206,000 and to place their records on Mass Storage.

The variables are read in from raw data and the first two pages of Table XXVI, show the variables locations and their descriptions. This lists all the variables available in the cohort even though only a smaal number of them were actually used for the present analysis.

The reader is referred to the SAS "User's Guide" for an explanation of the procedures involved. Comments throughout the listing explain some of the program steps.

2. SAS Program to Perform Regressions

The following two tables give listings for two SAS programs.

The first (Table XXVII) is the one used perform stepwise regression on the whole set of predictors so as to select out those which the more useful. The reader will note that after the SH data is read in, it is sorted by race which facilitates the subsequent regressions which are performed by race.

The second (Table XXVIII) uses the output from the previous analyses to construct regression models for relevant race/sex groupings. The data is split, for each

grouping, into two random samples used as the basis for cross-validation. Predictor and criterion scores, after validation are then output, in raw data form, to mass storage sc as to be available for later FORTRAN analysis in estimating cut-offs.

3. FCRTRAN Program used to Estimate Cut-offs on Fredictors

The raw output produced in the preceding SAS run is used as input data for this program (Table XXIX) along with the appropriate utility values.

The program is written so as to be run on one of the Navy Postgraduate School's terminals which is linked to a Tecktronics plotter, although it could easily be rewritten for a different plotting device. The plotting is performed using subroutines of the School's DISSPLA computer package.

TABLE XXVI

Listing of SAS Set-Up Program: Cohort to File //NESBITTS JOB (2501,0171), 'NESBITT ', CLASS=K
//*MAIN ORG=NFGVM1.2501P
// EXEC PGM=IFFBR14
//DD1 DD DISP=(OLD,DELETE),DSN=MSS.S2501.NRATESH
// EXEC SAS PHB1. PHB1. PPHB1. 200004 ASVABAR ASVABMK ASVABMC AS V A BW K
AS V A BS P
AS V A BE I
AS V A BG S
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ENTRY AGE = AGE CF INDIVIDUAL AT TIME OF ENTRY
EFCCFDID= RECORD ID--EXAM SCORE, DEP, ACTIVE
HYEC = HIGHEST YEAR OF EDUCATION
SEX = (1) MALE, (2) FEMALE
RACE = (1) WHITE, (2) BLACK, (3) OTHER
MRTLDEND=MARITAL STATUS/DEPENDENTS
TESTFORM=TEST FORM/ECFA, ASVAB, AFWST, AFOT, OSB
AFCTPCNT=AFOT PERCENTILE (OR EQUIVALENT)
AFOTGRPS=AFOT GROUPS (5,4C,4B,4A,3B,3A,2,1)
ASVAEGI = ASVAE APTITUDE AREA SCORE--SUBSCALE
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PRRTARRY = FRESENT RATE (ABBR.)

EXAMBATE = YAMINATION RATE CODE

EXHTARRY = EXAMINATION RATE CODE

ZISTARRY = ZAMINATION RATE (ABBR.)

TCILRAM = TOTAL RAW SCORE

SIDNAYY = STANDARDIZED NAVY SCORE

FISCALE = FROCES CODE

ALIPROBE = FROCES CODE

ALIPROBE = FROCES CODE

ALIPROBE = FROM SEC SCORE

FINLHULT = CANDILATE'S FINAL MULTIPLE

FOR THE FROM SEC SCORE

ARIFACTER = CHANGE OF RATE INDICATOR

ARIFACTER = CHANGE OF RATE INDICATOR

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IF DMDCRATE='SH' OR PRRTABRV='SH' OR
RCPGSCRT='2490' OR EXAMRATE='2490';
PFCC FREQ;
TABLES DMDCRATE PRRTABRV RCPGSCRT
AUTHRATE PRESRATE EXAMPATE;
TITLE ATTEMPT AT FORMING AN SH FILE;

/*

TABLE XXVII

Program Listing of Stepwise Regression for SH

```
//NESBITTS JOE (2501,0171), 'NESBITT ', CLASS=C
//*MAIN LINES=(25), ORG=NPGVM1.2501P
// EXEC SAS
//SAS.WORK DD SPACE= (CYL, (25,50)),
DISP= (NEW, DELETE, DELETE)
//FILEIN_CD_ CISP=SHR, DSN=MSS. S2501.NRATESH
 //SYSIN DD *
OFTICNS LS=80 NOCENTER NODATE;
 DATA; SET FILEIN. NRATESH;
       NUMBER OF YEARS OF EDUCATION IS CONVERTED FROM ITS DMDC ORDINAL CODING (1-13) TO A numeric.
IF HYEC=1 THEN CHYEC=3.5; IF HYEC=2 THEN CHYEC=8; IF HYEC=3 THEN CHYEC=9; IF HYEC=4 THEN CHYEC=10; IF HYEC=5 THEN CHYEC=11; IF HYEC=6 THEN CHYEC=12; IF HYEC=7 THEN CHYEC=13: IF HYEC=8 THEN CHYEC=14; IF HYEC=9 THEN CHYEC=15; IF HYEC=10 THEN CHYEC=16; IF HYEC=11 THEN CHYEC=18; IF HYEC=12 THEN CHYEC=20; IF HYEC=13 THEN CHYEC=11.5;
            VALIDITY VALUE SCREENS
IF ENTRYAGE GE 35 THEN ENTRYAGE=35;
IF ENTRYAGE >= 17; IF TOTPROMO<=5;
IF SCHLCODE='A' THEN SCHLCODE='1';
ELSE SCHLCODE='0': NUSCHCDE=SCHLCODE+0;
NUATTRIT=ATTRITCD+0; IF NUATTRIT=2 THEN
NUATTRIT=1; ELSE NUATTRIT=0;
NUHTFAY=HYPAYGRD+0;
NUHYFAY=HYPAYGRD+0;
NUHYFAY=HYPAYGRD+0:
YEAR=SUBSTR(LNGTHSRV,1,2);
MCNTH=SUBSTR(LNGTHSRV,3,2);
YEARS=YEAR+0; MCNTHS=MONTH+0;
LCSMNTHS=YEAR$*12+MONTHS:
IF MCNTHS >= 6 THEN YEAR$=YEAR$+1;
ONEYEAR=0:IF LCSMNTH$ >= 12 THEN ONEYEAR=1:
TWCYEAR=0:IF LCSMNTH$ >= 24 THEN TWOYEAR=1:
THRYEAR=0:IF LCSMNTH$ >= 36 THEN THRYEAR=1:
FORYEAR=0:IF LCSMNTH$ >= 48 THEN PORYEAR=1:
1AEEL LOSMNTH$=TOTAL LOS IN MONTH$ (NUMERIC);
     the highest paygrade is determined
IF FILEFLG1=8209 THEN PAYGRADE=PAYGRDE1;
IF FILEFLG1 NE 8209 THEN PAYGRADE=PAYGRDE3;
IF PAYGRADE>=2 THEN NCAYSE22=NDAYSE2:ELSE NDAYSE22=0;
IF PAYGRADE>=3 THEN NDAYSE23=NDAYSE3:ELSE NCAYSE23=0;
IF PAYGRADE>=4 THEN NCAYSE24=NDAYSE4;ELSE NDAYSE24=0;
IF NCAYSE22=0 THEN NCAYSE22=1275;
IF NCAYSE23=0 THEN NCAYSE22=1461;
IF NCAYSE23=0 THEN NDAYSE23=1461;
IF NCAYSE24=0 THEN NDAYSE24=1750;
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TIMEE4 = 0:

TIMEE2 = (ROUND (((NDAYSE22)/(365/12)

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TIMEE4 = (ROUND (((NDAYSE24)/(365/12)
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IF FILEFLG1=8209 THEN ELIGREUP=1;
IF ((FILEFLG1 NE 8209) AND (ISC3 GT 0) AND
(ELGREUP3 EQ 1)) THEN ELIGREUP=1; ELSE ELIGREUP=0;
            defines "ACHIEVED E-4", IN JOINT CONSIDERATION OF THE D.M.D.C. AND
                                                                                                                                                                                                                       N. H. R. C.
              ((PAYGRADE GE 4) AND (HYPAYGRD GE 4)) THEN ACHVDE4=1
((PAYGRADE LT 4) OR (HYPAYGRD LT 4)) THEN ACHVDE4=0;
                                                                                                                                                                                                                                                       THEN ACHYDE4= 1;
             defines "RATED" VERSUS "NOT-RATED".
 IF ((DMDCRATE NE '.') AND (DMDCRATE NE AND (SERVACCS EQ 2) AND (SERVICE1 EQ 2) ((FAYGRADE GE 4) AND (HYPAYGRD GE 4))) FATED=1; ELSE RATED=0;
            eliminates invalid asvab subscale scores [
             ASVABGI<=15:IF ASVABNO<=50:IF
ASVABWK<=30:IF ASVABAR<=20:IF
ASVABMK<=20:IF ASVABFI<=30:IF
ASVABGS<=20:IF ASVABSI<=20:IF
ASVABCM<=30:IF ASVABCA<=20:IF
ASVABCC<=30:
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IF (RCPGSCRT='2490' AND EXAMRATE='2490'
AND EMDCRATE='SH') THEN ENTRYGRP=1:
IF (RCPGSCRT='2490' AND EXAMRATE='2490'
AND EMDCRATE NE 'SH') THEN ENTRYGRP=2:
IF (RCPGSCRT='2490' AND EXAMRATE NE '2490'
AND DMDCRATE='SH') THEN ENTRYGRP=3:
IF (RCPGSCRT='2490' AND EXAMRATE NE '2490'
AND DMDCRATE NE 'SH') THEN ENTRYGRP=4:
IF (RCPGSCRT NE '2490' AND EXAMRATE='2490'
AND DMDCRATE='SH') THEN ENTRYGRP=5:
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IF (RCPGSCRT NE '2490' AND EXAMRATE='2490'
AND DMDCRATE NE '2490' AND EXAMRATE NE '2490'
AND DMDCRATE NE '2490' AND EXAMRATE NE '2490'
AND DMDCRATE NE '3490' AND EXAMRATE NE '2490'
AND DMDCRATE NE '3490' AND EXAMRATE NE '2490'
AND DMDCRATE='SH') THEN ENTRYGRP=7:
                   NEW VARIABLE "DEPNDNTS"
  IF MRTLDPND=10 THEN DEFNDNTS=0:
IF MRTLDPND GT 10 THEN DEPNDNTS=1:
  GOCDGUY = 20
  IF LCSMNTHS < 48 THEN GOODGUY =10;
IF LCSMNTHS >= 48 AND RATED=1
AND ELIGREUP=1 THEN GOODGUY =30;
EALGUY=30:
IF LOSMNTHS
ISC3>=60)
IF LOSMNTHS
IF LOSMNTHS
IF LOSMNTHS
                                                                 S < 48 AND (ISC3 <=87 AND THEN BADGUY =10;

< 48 AND TOTIAWOL > 0 CONTROL > 0 
                                                                                                                                                                                                                    THEN
THEN
THEN
THEN
                                                                                                                                                                                                                                                BADGUY = 10
                                                                                                                                                                                                                                                BADGUY = 10
                  LCSM NTHS
LCSM NTHS
                                                                                                                                                                                                                                              BADGUY
BADGUY
DGUY =20
                                                                                                                                                                                                                                                                                             = 10
                                                                                                                                                                                                                                                                                             = 10
 Ī
                 LCSMNTHS
                                                                                                                                                                                                  THEN
                                                                                                                                                                                                                                BADGUY
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IF LCSMNTHS >= 48 AND TOTLDEMO > 0 THEN BADGUY = 20: IF LCSMNTHS >= 48 AND ELIGREUP = 0 THEN BADGUY = 20:
IF LCSMNTHS >= 48 AND TOTLDEMO > 0 THEN BADGUY = 20:
LABEL
LABEL
LEEN NTS >= 48 AND TOTLDEMO > 0 THEN BADGUY = 20:
LABEL
LEEN NTS >= 48 AND ELIGREUP = 0 THEN BADGUY = 20:
LABEL
LEEN NTS SINGLE, NO KIDS OVTHERS 1
CHYBC = CONVERTED NUMBER OF YEARS OF EDUCATION
NUMBER OF YEARS OF EDUCATION
NUMBER OF YEARS OF EDUCATION
NUMBER OF PAYGRADE ATTAINED
NUMBER OF PAYGRADE ATTAINED
NUMBER OF DAYS TO E-0
NUATTRIT = NHRC FILE--ATTRITION CODES
NUNCKOCK = NHRC -NOT RECOMMENDED FOR RE-ENLISTMENT
NLAYSE2 = SCREENED NUMBER OF DAYS TO E-0
NUATTRIT = SCREENED NUMBER OF DAYS TO E-0
NUATSE2 = SCREENED NUMBER OF DAYS TO E-0
NUA
                 LABEL
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TABLE XXVIII

Program Listing to Output & Cross Validate

```
//NESEITTS JOB (2501, C171), 'NESBITT ', CLASS=C
//*MAIN LINES= (25), ORG=NFGVM1.2501P
// EXEC SAS
         /*MAIN LINES= (25), ORG=NFGVM1.2501P

/*EXEC SAS

//SAS.WORK DD SPACE= (CYL, (25,50))

// DISP=(NEW, DELETE, DELETE)

//FILFIN DD DISP=SHR, DSN=MSS.S2501.NRATESH

//FT31F001 DD UNIT=3330V, MSVGP=PUB4Z,

DLSP=(NEW, CATLG), DSN=MSS.S2501.SHLMN,

DCB=(RECFM=FB, LRECL=80, BLKSIZE=3120)

//FT32F001 DD UNIT=3330V, MSVGP=PUB4Z,

DISP=(NEW, CATLG), DSN=MSS.S2501.SHLWH,

DCB=(RECFM=FB, LRECL=80, BLKSIZE=3120)

//FT33F001 DD UNIT=3330V, MSVGP=PUB4Z,

DISP=(NEW, CATLG), DSN=MSS.S2501.SHLWH,

DCB=(RECFM=FB, LRECL=80, BLKSIZE=3120)

//FT34F001 DD UNIT=3330V, MSVGP=PUB4Z,

DISP=(NEW, CATLG), DSN=MSS.S2501.SHLBL,

DCB=(RECFM=FB, LRECL=80, BLKSIZE=3120)

//FT34F001 DD UNIT=3330V, MSVGP=PUB4Z,

DISP=(NZW, CATLG), DSN=MSS.S2501.SHLOT,

DISP=(NZW, CATLG), DSN=MSS.S2501.SHLOT,

DCB=(RECFM=FB, LRECL=80, BLKSIZE=3120)
FIT34 TO 0 TO DISTIBLE TO SAMPLE TO SET OF TOTAL LOS MITHAT: MCD EL

LISMITHS TO A STABBLE AS VABBLE AS VABBLE TOOLE

LISMITHS TO A STABBLE AS VABBLE AS VABBLE TO SET OF SET OF
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```
FRCC SORT;

TITLE SORTING BY PRDICTOR;
TATA STORE; SET TABLE; FILE FT31F001;
PUT (FORYEAR PRDICTOR THREAR RACE) (10.7)
THILE SORTING BY PROLUCTOR;

ITTIE SORTING BY PROLUCTOR FOR ALL SH MEN;

INTIA STORE;

INTIA TRIALCAT;

INTIA TRIALCAT;

INTIA TRIALCAT;

INTIA STORE;

INTI
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```
MRISTATI ASVABCC AFOTGRPS;
LATA TABLE;SET TWOPRED;
PRDICTOR=(LOSM 1HAT+LOSM2HAT)/2;
PRCC SORT;
ENTILE SORTING EY PRDICTOR FOR BLACK SH MEN;
DATA STORE;SET TABLE;FILE FT33 P301;
PUT (PORYEAR PREDICTOR THRYEAR RACE) (10.7);
DATA TRIALDAT; SET DATA 1;
FRACE=3;* SELECTS OTHERS SHS;
DATA CERFPNTZ; SET TRIALDAT;
DATA CERIV8; SET DIFFRNTZ; FRANDNUM1>1;
DATA LERIV8; SET DIFFRNTZ; FRANDNUM1>1;
DATA LERIV8; SET DIFFRNTZ; FRANDNUM1=1;
DATA VALID8; SET DIFFRNTZ; FRANDNUM1=1;
DATA SAVABGI ASVABAR ASVABGS ASVABMC ENTRYAGE

LCSMNTHS=ASVAEGI ASVABAR ASVABGS ASVABMC ENTRYAGE

LCSMNTHS=ASVAEGI ASVABAR ASVABGS ASVABMC ENTRYAGE

TILLE MODEL OTHERS SH SFOR LOS SAMPLE TWO;
PRCC PEG DATA=VALID8 SIMPLE OUTEST=B01; DOM 2HAT; MCD PL
LCSMNTHS DATA=VALID8 SIMPLE OUTEST=B01; DATA HAT; MCD PL
LCSMNTHS BAVABGI ASVABBAR ASVABGS ASVABMC ENTRYAGE

TITLE MODEL OTHERS SHS FOR LOS SAMPLE TWO;
PRCC SCORE OUT=E01PRED TYPE=OLS SCORE=B01

ASVABGO ASVABMC ENTRY AGE ASVABAD;
PRCC SCORE OUT=E01PRED TYPE=OLS SCORE=B02

LATA=LIRIV8 PREDICT; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;
FRCC CORE DATA=E01PRED; VAR

TITLE OTHERS SECOND VALIDITY COEFFICIENT;
FRCC CORE DATA=E01PRED; VAR

TITLE OTHERS SECOND VALIDITY COEFFICIENT;
FRCC CORE DATA=E01PRED; VAR

ASVABGS ASVABMC ENTRY AGE ASVABAD;

TITLE OTHERS SECOND VALIDITY COEFFICIENT;
FRCC CORE OUT=HOPRED;
PROTOCOCCUM THAT PREDICT; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

TITLE OTHERS SECOND VALIDITY COEFFICIENT;
FRCC CORE OUT=FOICT; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

FRCC SCORE OUT=TWOPRED;
PROTOCOCCUM THAT AND ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

FRCC SCORE OUT=FOICT; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

FRCC SCORE OUT=FOICT; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

FRCC SCORE OUT=FOICT T; VAR ASVABGI ASVABAR
ASVABGS ASVABMC ENTRY AGE ASVABAD;

FRCC SCORE OUT=FOICT T; VAR ASVABGI ASVABAR
ASVABGS ASV
```

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TABLE XXIX

Program Listing of FORTRAN Program for Cut-offs

```
REAL PASS (4000) .SCORE (3,4000) .PASST (4000) .

RETURN (3,4000) .TOTRET (3,3) .NPLOT (3) .

SUCALL (3) .RACE (4000) .SRALL (3) .UTILSP (2) .

UTILSF (2) .RACE (4000) .TOTSC (3,3) .

S (4000) .R (4000) .TMS (3) .TMR (3) ...

CATA UTILSP .- 9244 . .18488 .0 ./

IATA UTILSF / 14495 .0 .- 14495 ./

SINGLE = 3 .

C = 55 .0

E = 25 .0

E = 25 .0

E = 1

FASSA = 0 .0
                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                               00 10
                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                                           *
                                   I = 1

FASSA = 0.0

FASSB = 0.0

FAILA = 0.0

FAILB = 0.0

SCORE (J, I) = 30.0

FETURN (J, I) = 0.0

RETURN (J, I + 1) = 0.0

FACE (I) = 0.0

FORMAT (4P10.7)

I = I + 1

READ(J, 10, END=50) PASS (I), SCORE (J, I),

PASST (I), RACE (I)
20 I=I+1
READ(J, 10, END=50) PASS(I), SCORE(J, I),

*

*PASSA = PASSA + PASS(I)

50 CCNTINUE
NFLOT(J) = I - 1
GTOTAL = I - 2
FAILA = GTOTAI - PASSA
X = PASSA * UTILSP(2) + PASSB * UTILSP(1)
X = X + FAILA * UTILSF(2) + FAILB * UTILSF(1)
BASSUT = X/GTCTAL
N = NPLOT(J)

109 FORMAT(I4, F10.3, F10.2, I8)
109 FORMAT(I4, F10.3, F10.2, I8)
109 FORMAT(I4, F10.3, F10.2, I8)
11 FAILB = FAILB + 1
FAILA = FAILA - 1
GO TO 33

32 FASSB = PASSB + 1
33 X = FASSA * UTILSP(2) + PASSB * UTILSP(1)
X = X / GTOTAL
IF (BASEUT - NE.0.0) GO TO 77
BASEUT = X
IF (BASEUT - NE.0.0) GO TO 77
BASEUT = X
IF (BASEUT - NE.0.0) GO TO 77
BETURN(J, I + 1) = 100 * (X - BASEUT) / (BASEUT)
IF (BASEUT - STORE)

77 RETURN(J, I + 1) = - RETURN(J, I + 1)
1F (BASEUT - STORE)
25 IF (I.EO.2) GC TO 34
1F (BASEUT - STORE)
34 SCCMAX(J, I) = SCORE(J, I)
SRALL(J) = (FASSA) / (PASSA+FAILA) * 100
1F (SRALL(J) = (GASSA) / (PASSA+FAILA) * 100
1F (SRALL(J) = (GASSA) / (PASSA+FAILA) * 100
1SUCALL(J) = (GASSA) / (PASSA+FAILA) * 100
125
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06 10
06 20
06 30
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                                                                                                                                                                                                                                                                                                                                                                                                                                                                                            1040
1050
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1080
1100
11120
1130
                                                                                                                                        (S,R,N,C)
(IMS,TMR,3,0)
(1.09)
(S ("M+AXIMUM /R+ETURN: ',19,
TOTSC (J,3),TOTRET (J,3))
                                         IF (J.EQ.1)
IF (J.EQ.2)
CALL RLMESS
GO TO 130
CALL RLMESS
GC TO 130
CALL RLMESS
CALL RLMESS
                                                                                                                                                          GO TO 110
GO TO 120
('O+THERS',7,'ABUT','ABUT')
                                                                                                                                                                                                                                                                                                                                                                                                                                                                                           1140
                                                                                                                                                               ('B+LACKS', 7, 'ABUT', 'ABUT')
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11
                                                                                                                                                             ('W+HITES',7,'ABUT','ABUT')

(SCOMAX(J,1),2, TOTRET(J,3)-2.)

TOTSC(J,3)+2. TOTRET(J,3)-2.)

(',',1,'ABUT','ABUT','ABUT')

('%',1,'ABUT','ABUT','ABUT')
                                            CALL RLMESS
CALL RLMESS
CALL RLMESS
CONTINUE
          200
                                           CUNTINGE

C=5.0

E=-5.0

CAIL RLMESS ('SELECTION RATIOS:', 17,55.0,D)

CAIL RLMESS ('PERCENTAGE SUCCESSFUL:',

22,55.0,E)
                                            DO 300
D=C-2.0
E=E-2.0
                                                                                                               J=1,3
                                                                            (J.EQ.1) GO TO 310
```

```
IF (J.EQ.2) GO TO 320

CAIL RLMESS ('O+THER /G+ROUP: ',18,55.0,D) 1350

CAIL RLMESS ('S',1,'ABUT','ABUT') 18,55.0,E) 1360

CAIL RLMESS ('X',1,'ABUT','ABUT') 18,55.0,E) 1380

CAIL REALNO (SUCALL(J),1,'ABUT','ABUT') 1400

CAIL RLMESS ('M',1,'ABUT','ABUT') 1440

CAIL RLMESS ('B+LACK /G+ROUP:',18,55.0,D) 1420

CAIL RLMESS ('B+LACK /G+ROUP:',18,55.0,D) 1440

CAIL RLMESS ('B+LACK /G+ROUP:',18,55.0,E) 1450

CAIL RLMESS ('B+LACK /G+ROUP:',18,55.0,E) 1460

CAIL RLMESS ('B+LACK /G+ROUP:',18,55.0,E) 1450

CAIL RLMESS ('M',1,'ABUT','ABUT') 1460

CAIL RLMESS ('M',1,'ABUT','ABUT') 1480

CAIL RLMESS ('M',1,'ABUT','ABUT') 1500

CAIL RLMESS ('M',1,'ABUT','ABUT') 1500

CAIL RLMESS ('M',1,'ABUT','ABUT') 1550

CAIL RLMESS ('M',1,'ABUT','ABUT') 1550
```

APPENDIX B STEPWISE REGRESSION BESULTS FOR RATINGS

This appendix gives the results of the stepwise regressions. They are presented in the form of rating specific tables. As these tables were constructed directly from computer output the variables all have the SAS names that were used in the analysis. Therefore, interested readers are directed to the relevant sections of the preceding SAS program listings to determine exactly which variables are listed. In all cases the variables and their coefficient signs correspond to the tables in the body of the thesis which lists the stepwise regression results.

TABLE XXX
All SH Stepwise Regression Results

	SH Whites on Length of	<u>Service</u>
REGRESSION ERROR TCTAL	SH Whites on Length of DF SUM CF SUU 8 16236.804 11. 1297 237637.401 1305 253874.205 B VALUE STD E 26.67C1644	<u>Service</u> 58 0.0001
	B VALUE STD E 26.6701644	F PROB>F
INTERCEPT AFTOTGEPS ASVAEAL ASVAECA ENTERPAYG MRISTAT 1 SCREEN CEPNDNTS	-0.2688282 0.1360 3. -0.1703922 0.0978 3.	99 0.0145 90 0.0484 03 0.0818 87 0.0903 42 0.0357 89 0.0001 73 0.0001 20 0.0739
REGRESSION ERROR TCTAL	-0.15 ± 50 4 9 0.0917 2.170 9 373 4 0.8 128 4.5.26 128 4.5.26 128 128 128 128 128 128 128 128 128 128	EUY F PROB>F 22 0.0001
INTERCEPT ASVAEMK ASVAECE -	1.10711196 0.08789585 0.0561 2.	F PROB>F
ASVABCE - CHYEC MFISTAT 1	0.08789585 0.0561 2. 0.14122207 0.0493 8. 1.36078242 0.2466 30. 1.09421802 0.4740 5.	45 0.1174 18 0.0043 43 0.0001 33 0.0211
D.T	SH Whites on BADGU 7 4135.7682 12. 1298 60204.2011 1305 64339.9693	Y F PROB>F 74 0.0001
REGRESSION ERROR TCTAL	SH Whites on BADGU DF SUM CF SOU 7 4135.7682 12. 1298 60204.2011 1305 64339.9693	74 0.0001
	1305 64339.9693 B VALUE STD E 1.74204431 0.03938113 0.0225 3.	F PROB>F
INTERCEPT - ASVABNO ASVABUK ASVABCE CHYEC ENTRPAYG MRISTAT 1 SCREEN		0.0816 47 0.0037 45 0.0352 58 0.0001 80 0.0287 61 0.0053 56 0.0186
	SH Blacks on Length of DF SUM OF SUM	<u>Servica</u> F PRO3>F 26 0.0001
REGRESSION ERROR IOTAL	SH Blacks on Length of DF SUM OF SUM	25 0.0001 F PROB>F
INTERCEPT AFOTGERS	32.7376235 -1.9561139 0.4563 18.	
INTERCEPT AFOTGFPS ASVABCA ASVAECC MRISTAT 1 SCREEN	6.2469389 1.1323 30. 0.2155208 0.0872 6.	37 0.0001 93 0.0087 35 0.0211 43 0.0001 10 0.0138
850350070	DF Blacks on GOODG DF SUM OF SOU 4 621.0247 3.	TF PROB>F
REGRESSION ERROR TCTAL	553 25714.4591 557 26335 4838	
INTERCEPT	11.4082688	F PROB>F
APQTGRPS	-0.66 12232 0.2428 7.	41 0.0067

ASVAECC MRISIAT1 SCREEN		92502 35432 55462	0.0590 0.6035 0.0460	2.83 3.59 3.45	0.0931 0.0586 0.0639
REGRESSION ERROR TCTAL	D 5 5 5 5	SH Blac F SUE 4 90 3 1939 7 1929	ks on BA CF 500 0.1514 9.6692 9.8207 STD E	6.76	PROB>F 0.0001
INTERCEPT ASVAEAR ASVAECC ENTRYAGE CHYEC	11.19 -0.24 0.10 -0.19 1.16	SH Blac 9999 1929 1929 1929 1929 1929 19404 194049 19404 194	STD E 0.0778 0.0497 0.1156 0.3167	F 10.23 4.38 2.84 13.60	PROB>F 0.0015 0.0369 0.0926 0.0002
REGRESSION ERROR TCTAL	SH Ot	hers cn	Length CF 500 6.5314 4.4685 1.0000 STD E	<u>of Serv</u> 11.36 F	FROB>F
INTERCEPT ASVABGI ASVAFAD ASVAFAR ASVAEMC ASVAEGS ENTRYAGE	54.99 -1.17 -0.33 -C.87 -0.71	6 1887 1887 2751 2751 34110 2115 2115 2115 2115 2115 2115 2115 2	0.3280 0.1989 0.2952 0.3736 0.3674 0.3009	12.74 2.77 8.88 3.62 10.11 4.17	0.0005 0.0984 0.0034 0.0591 0.0018 0.0430
REGRESSION ERROR TCTAL	15 15	SH Othe F SUM 5 105 0 529		00DGUY 5.95	PROB>F 0.0001 PROB>F
INTERCEPT AFOIPCNT ASVAEEI ASVAEAI ASVAECC SCREEN	8.680 -0.094 -0.163 -0.288 -0.387 0.146	VALUE 173217 86851 84685 115095 197354 137842	0.0286 0.1001 0.1548 0.1097 0.0801	10.93 2.68 3.46 12.49 3.33	0.0012 0.1039 0.0647 0.0005 0.3698
REGRESSION ERROR ICTAL	D	SH Othe		<u>DGUY</u> 4.89	PROB>F 0.0004 PROB>F
INTERCEPT AFCTFCNT AFCTGRPS ASVAEAI ASVAECC CHYEC	6.770 -0.161 1.315 -0.295 0.235	959 959 950 960 960 960 960 960 960 960 960 960 96	0.0584 0.7261 0.1288 0.0952 0.4234	7.67 3.28 4.90 6.29 5.37	0.0063 0.0720 0.0284 0.0147 0.0218

TABLE XXXI
All PN Stepwise Regression Results

<u> </u>	white Men on Length of Serv	ice
REGRESSION ERROR TOTAL	White Men on Length of Serving 7 50M OF SOU F 9 23114.535 14.78 1216 211247.111 1225 234361.647 B VALUE STD E F	5.0001
~ u ~ D D C D D D	B VALUE SIDE F	PROB>F
INTERCEPT ASVAEWK ASVAEER ASVAEEII ASVAECE ENTRYAGE CHYEC MRISTAT 1 SCREEN	1216 211247.111 1225 234361.647 B VALUE STD E F 30.6886800 -0.3907675 0.0918 18.10 -0.2914101 0.1277 5.20 0.2812350 0.1048 7.19 -0.2984347 0.1175 6.44 0.1368159 0.0849 2.59 0.3694307 0.1831 4.07 -1.0215411 0.4358 5.49 7.6648584 0.7969 92.49 0.2812599 0.0781 12.96	0.0001 0.0227 0.0074 0.0113 0.1076 0.0439 0.0193 0.0001
	PN White Men on GOODGUY DF 500 CF 500 5.33	
REGRESSION ERROR TCTAL	PN White Men on GOODGUY 8 2013.1708 5.33 1217 57483.8112 1225 59496.9820 B VALUE STD E F	PROB>F 0.0001
	1225 59496.9820 B VALUE SID E F 10.4760049	PROB>F
INTERCEFT ASVABAR ASVABBCE ASVABCE CHYECAT MRISTAT SCREEN DEPNDNTS	PN White Men on GOODGUY 8 2013.1708 5.33 1217 57483.8112 1225 59496.9820 B VALUE STD E 10.4760049 -0.1122275 0.0443 6.41 -0.2187153 C.0682 10.28 0.0974479 0.0687 2.34 0.1244157 0.0424 8.58 0.1244157 0.0424 8.58 -0.3040941 0.1929 2.48 0.9357586 0.4190 4.99 0.1642128 0.0406 16.33 1.5133497 0.8221 3.39	0.0115 0.0014 0.1267 0.00354 0.1257 0.0059
CHYEC MRISTAT1	-0.3040941 0.1929 2.48 0.9357586 0.4190 4.99	0.0257
SCREEN DEPNDNTS	0. 16 42128 0.0406 16.33 1.5133497 0.8221 3.39	5.5659
	PN White Men on BADGUY	PROBYE
REGRESSION ERRCR TOTAL	1221 33550.0625 1225 34176.2642	PROB>F 0.0002
	B VALUÉ STD E F 12-9046054	PROB>F
INTERCEPT AFCTGRPS ASVABEI SCREEN DEPNDNTS	12.9046054 -0.5259052 0.1598 10.83 0.0801835 0.0373 4.60 0.1092251 0.0268 16.54 1.1273538 0.6069 3.45	0.0010 0.0321 0.0001 0.0635
SCREEN DEPNDNTS	0.1092251 0.0268 16.54 1.1273538 0.6069 3.45	0.0635
_	N_Black Men on Length of Seri	Zics PROB>F 0.0001
REGRESSION	N Black Men on Length of Serior DF 500 OF 500 8.31 276 37349.7756 282 44100.6431	5.0001
REGRESSION ERRCR TOTAL	276 37349.7756 282 44100.6431 B VALUE STD E F 48.8556818 -0.1751397 0.0417 17.17 -0.1314724 0.0864 2.31 -0.3337261 0.1623 4.23 -2.4543496 0.5735 18.31 5.8937313 1.4492 16.54 0.4812862 0.1509 10.17	P RO B > F
INTERCEPT AFÇTECNT	48.8555818 -0.17:1397 0.0417 17.17	0.0001
ASVABNO ASVAECC CHYEC	-0.1314724 0.0864 2.31 -0.3337261 0.1623 4.23	0.1295
MRISTATI	-0.17:1397 0.0864 2.31 -0.13:14724 0.0864 2.31 -0.33:37261 0.1623 4.23 -2.45:43496 0.5735 18.31 5.89:37313 1.4492 16.54 0.48:12862 0.1509 10.17	0.1295 0.0408 0.0001 0.0001 0.0016
SCREEN	0.4812862 0.1509 10.17	9.0016
	PN Elack Men on GOODGUY DF SUM OF SOU 7 950.4489 3.68	PROB>F 0.0009
REGRESSION EFROR TOTAL	PN Elack Men on GOODGUY DF 50M OF 50U 7 950.4489 3.68 275 10156.2648 282 11106.7137	0.000

```
B VALUE
24.95382658
0.3102658
-0.1398792
0.1769179
0.1726309
-0.3085155
-0.1521711
-G.5306986
                                                                                                                                                                                  STD E
                                                                                                                                                                                                                                                                            F
                                                                                                                                                                                                                                                                                                           PROB>F
INTERCEPT
                                                                                                                                                                           C.1504
0.0472
0.1037
0.0956
0.1207
0.0810
0.2566
                                                                                                                                                                                                                                                    4.26
8.77
2.96
5.53
4.27
                                                                                                                                                                                                                                                                                                         0.0401
0.0033
0.0894
0.0721
0.0112
 ASVAEGI
ASVAENO
ASVAEMK
ASVABEI
 ASVAEST
ASVABCE
                                                                                                                                                                                                                                                                                                            0.0396
  CHYEC
                                                                               PN Black Men on BADGUY
DF SUM CF SQU F
4 390.1637 5.31
278 5106.6559
282 5496.8197
B VALUE STD E F
27.3597073
-0.2845607 0.0866 10.80
0.2497785 0.0890 7.87
-0.3561344 0.1018 12.22
0.5638719 0.3534 2.54
                                                                                                                                                                                                                                                                                                           PROE>F
0.0004
REGRESSION
ERROR
 TCTAL
                                                                                                                                                                                                                                                                                                           PROB>F
INTERCEPT
ASVABSI
ASVAEAI
ENTRYAGE
ENTRPAYG
                                                                                                                                                                                                                                                                                                          0.0011
                                                                PN Other Mer on Length of Service
DF SUM OF SCU
N 3 3425.1745 9.05 0.0001
87 10976.1221
90 14401.2967
B VALUE STD E F PROB>F
37.3172602
-0.9104034 0.3173 8.23 0.0052
1.8625773 0.8067 5.33 0.0233
-12.7447520 3.9740 10.28 0.0019
 REGRESSION
ERRCR
  TCTAL
INTERCEPT
ASVABGI
CHYEC
 DEFNDNTS
                                                                       GOODGUY
                                                                                                                                                                                                                                                                                                           PROB>F
0.0005
                                                                                                                                                                                                                                                      4.99
REGRESSION ERRCR
 TCTAL
                                                                                                                                                                                                                                                                             F
                                                                                                                                                                                                                                                                                                           PROB>F
INTERCEPT
AFCTPONT
ASVABAR
                                                                                                                                                                                                                                               9.28
2.18
5.68
10.57
5.55
                                                                                                                                                                                                                                                                                                          0.0031
0.1437
0.0194
0.0016
0.0208
ASVAECC
CHYEC
CEFNENTS
PN Other Men of Street 
                                                                                                                                                                                                                 on BADGUY
                                                                                                                                                                                                                                                                                                           PROB>F
0.0007
                                                                                                                                                                                                                                                        5.34
                                                                                                                                                                                                                                                                              7
                                                                                                                                                                                                                                                                                                           PROB>F
                                                                                                                                                                                                                                               5.16
8.44
10.89
4.48
                                                                                                                                                                                                                                                                                                         0.0257
0.0047
0.0014
0.0372
                                                          PN White women on Length of Service

DF SUM CF SUM F PROB>F

10N 6 7047.9736 6.31 0.0001

472 87924.6777

478 94972.6513

B VALUE STD E F PROB>F

95.2187469

-0.3924025 0.1608 5.95 0.0151

-0.3429857 0.1734 3.91 0.0485

-0.4423221 0.2133 4.30 0.0387

0.5541282 0.1689 10.75 0.0011
 REGRESSION
ERROR
  TCTAL
 INTERCEPT
ASVAEAR
ASVAEAR
ASVAEAR
ASVAECE
```

CHYEC ENTRPAYG	-2.7667462 0.8557 1.7744842 1.1070	10.45 2.57	0.0013 0.1096
REGRESSION ERRCE TOTAL	PN White Women on 3 971.0785 475 27338.7335 476 28309.8121	GOODGUY 5.62	PROB>F 0.0010
INTERCEPT	B VALUE STD E 29.1338691	F	PROB>F
ASVAEAD ASVAEWK ASVAEAR	-0.2337727 0.0893 -0.1576776 0.0935 -0.2607376 0.1169	6.85 2.84 4.97	0.0092 0.0924 0.0262
	PN White Women on	BADGUY	
REGRESSION ERROR TCTAL	5 576.7302 473 10619.5119 478 11196.2421	5.14	PROB>F 0.0002
INTERCEPT	B VALUE STD E 19.4431146	F	PROB>F
ASVABAD ASVABAR ASVABSP CHYEC DEPNONTS	-0. 12 12187	4.75 8.57 2.91 4.47 7.22	0.0299 0.0036 0.0887 0.0351 0.0075

TABLE XXXII
All AT Stepwise Regression Results

<u>A T</u>	White Men on Length of Servi	ce Ce
REGRESSION ERRCR TOTAL	White Men on Length of Service 12 37 330.665 29.55 3327 350277.988 3339 387 608.653 B VALUE STD E F	0.0001
INTERCEPT	B VALUE STD E F	PROB>F
INTERPREDICT AFCORDS AFCORDS AFCORDS ASVABBS ASVABBCA ASVABBCA ASVABBCA ASVABBCA ASVABBCA ASVABCA ASVA	3339 387608.653 B VALUE STD E F 49.7303064 0.0847810 0.0278 9.24 -0.8119479 0.5030 2.60 -0.0788686 0.0243 10.46 -0.1110189 0.0670 2.74 -0.1183909 0.0684 2.99 -0.1355446 0.0712 3.61 -0.2048483 0.0586 12.19 0.1205116 0.0453 7.07 -0.7428230 0.2858 6.75 2.1286892 0.1946 119.62 4.6121976 0.3735 152.47 0.0957458 0.0406 5.56	0.0024 0.1062 0.10977 0.09379 0.00079 0.00099 0.00001 0.00185
	AT White Men on GOODGUY DE SUN OF SOU 3.27	PROB>F 0.0202
REGRESSION ERRCR TOTAL	AT White Men on GOODGUY F 500 OF 500 A 56.839 3.27 3336 155283.609 3339 155740.449 B VALUE STD E F	0.0202
	B VALUE STD E F	FROB>F
INTERCEPT CHYEC ENTERAYG MRISIAII	3336 155 283.699 3339 155740.449 B VALUE STD E F 13.8263146 0.3664369 0.1663 5.40 -0.1868022 0.1270 2.16 0.5036740 0.2450 4.23	0.0202 0.1415 0.0399
	AT White Men or BADGUY	FROB>F
REGRESSION ERRCR TOTAL	6 4061.6007 32.17 3333 70130.5848	PROB>P 0.0001
	B VALUE SID E F	PROB>F
APOTGRPS ENTRYAGE	B VALUE SID E F 12.1536983 -0.3680884	0.0003
CHÝEC ENTEPAYG	0.4063458 0.1381 8.65 -0.9321898 0.0862 116.75	0.0033
INTERCEPT APOTGRES ENTRYAGE CHYEC ENTREAYG MRTSTATI SCREEN	TO 1536983 -0.3680884 0.1027 0.4063458 0.4063458 0.1027 0.4063458 0.1381 0.4063458 0.1381 0.4063458 0.1381 0.4063458 0.1669 0.4857506 0.1669 0.4857506 0.1669 0.4857506 0.1832364	0.0003 0.0541 0.0033 0.0001 0.0036 0.0001
<u>aI</u>		
REGRESSION ERRCR TOTAL	Elack Men on Length of Service 5 37 77. 9917 7.66 167 16480.0660 172 20258.0578 B VALUE STD E F 76.0373893 -0.6729970 0.2328 8.35	5.0001
INTERCEPT	5 VALUE STD E F	PROB>F
ASVAEGS ASVAESI	-0.6729970 0.2328 8.35 -0.4349189 0.2312 3.54	0.0044
ASVAECA CHYEC	-0.6729970 0.2328 8.35 -0.4349189 0.2312 3.54 0.3428155 0.1777 3.72 -1.7461861 7.7264 5.78 4.5120091 1.5433 8.55	0.0044 0.0617 0.0555 2.0173 0.0039
MRTSTAT1	AT Flack Mon on COORGHY	0.0039
REGRESSION	AT Flack Men on GOODGUY F SUM OF SOU 7.99	PROB>F 0.0001
ERROR TOTAL	167 5181.4687 172 6420.8092	
	B VALUE STD E F	PROB>F

```
INTERCEPT
AFCIFONT
AFOTGRES
ASVAEGS
                                     47. 2311662
0.2447246
-4.9159979
-0.2445383
0.1803334
-1.0599272
                                                                             0.0711
1.2330
0.1597
0.1012
0.4173
                                                                                                            11.82
15.89
2.34
3.17
                                                                                                                                        0.0007
                                                                                                                                       0.0001
0.1277
0.0768
0.0120
ASVAECA
CHYEC
                                                                                                               6.45
                                    AT Black Men 22

DF SUM OF 500

6 461.4027

166 2191.7764

172 2653.1791

B VALUE STD E

29.5616565

0.1124665 0.0452

-3.0880771 0.7941

0.3320927 0.1154

-0.1282669 0.0869

0.1147313 0.0618

-0.6578062 0.3131
                                                                                                    BADGUY
                                                                                                                                        PROB>F
REGRESSION ERRCR
                                                                                                               5.82
                                                                                                                                        0.0001
 TOTAL
                                                                                                                          F
                                                                                                                                       PROB>F
INTERCEPT
AFCTFCNT
AFCTGRPS
ASVAEAR
                                                                                                                                        0.0140
                                                                                                               6.17
                                                                                                            15.12
8.28
2.18
3.44
                                                                                                                                       0.0001
0.0045
0.1421
0.0653
0.0372
ASVABAK
ASVAEAI
                                                                                                               4.41
ENTRPAYG
                         AT Other Mer on Length
DF SUM OF SOU

ION 11 5054.3481
98 7557.6518
109 126 12.0000
B VALUE STD E

FT -4.67799655
0.60422952 0.2241
-0.99876644 0.2561
0.77440682 0.1924
E 1.40536754 0.4478
0.71935965 0.1641
                                                                                                     <u>of Servi</u>
                                                                                                                                    <u>.C.e</u>
PROB>F
REGRESSION ERRCE
                                                                                                                5.96
 TOTAL
                                                                                                                                        PROB>F
                                                                                                                          F
INTERCEPT
ASVABELI
ASVABAI
                                                                                                            7.27
15.21
16.20
9.85
19.22
                                                                                                                                       0.0083
0.0002
0.0001
0.0022
0.0001
ASVAECM
ENTRYAGE
SCREEN
                                AT Cther Men on G

1F 50 M OF 500

10 22 95 94 45

9 32 75 87 36

109 55 71 318 1

8 VALUE STD E

8 85 7501 16

-0.18 4255 71 0.050 4

0.51 4585 63 0.136 9

0.86 805 391 0.16 48

-0.330 352 20 0.1353

0.34 2139 80 0.168 9
                                                                                                  GOODGUY
                                                                                                                          F
                                                                                                                                        PROB>F
REGRESSION
ERROR
TOTAL
                                                                                                                6.94
                                                                                                                                        5.0001
                                                                                                                          F
                                                                                                                                        PROB>F
INTERCEPT
AFOIPONT
ASVAEAD
ASVAEWK
ASVAEEI
SCREEN
                                                                                                            13.35
14.11
27.72
5.96
9.86
                                                                                                                                        0.0004
                                                                                                                                        0.0003
                                                                                                                                        0.0164
                                    AT Other Men on BADGUY
DF SUM OF SUU
12 848.2633 6.55
97 1047.1911
109 1895.4545
B VALUE STD E
13.4256325
0.3560198 0.1186 9.00
-0.3764082 0.0825 20.77
-0.2365056 0.1077 4.81
                                                                                                                                        PROB>F
REGRESSION
ERRCR
                                                                                                               0.55
                                                                                                                                        0.0001
 TOTAL
                                                                                                                          F
                                                                                                                                        PROB>F
INTERCEPT
ASVAEMK
ASVABEI
                                                                                                            9.00
20.77
4.81
5.55
                                                                                                                                        0.0034
                                                                                                                                        0.0001
                                                                                                                                        0.0306
0.0205
 ASVAEAI
 SCREEN
                       AT white women on Length
DF SUM OF SQU
38 20.0507
235 44438.4450
239 48258.4958
B VALUE STD E
EPT 39.2507434
0.7791086 0.2424
0.3521168 0.2423
-0.5763499 0.2557
                                                                                                            <u>of</u>
                                                                                                                      <u>Servi</u>
                                                                                                                                        Ce
PROB>P
                                                                                                                5.05
 REGRESSION
                                                                                                                                        0.0006
 ERRCE
 TOTAL
                                                                                                                          F
                                                                                                                                        PROB>F
INTERCEPT
ASVABEI
ASVABOM
                                                                                                            10.33
2.11
5.08
                                                                                                                                        0.0015
                                                                                                                                       0.1476
 ASVAECE
```

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MRISTAT1	-5.1232380	1.7929	8.17	0.0047
REGRESSION ERRCE	239 110	OF SQU 45.1452 38.3938	<u>GOODGUY</u> 3.93	PROB>F J.C092
TOTAL	242 115 B VALUE 12.9788322	83.5390 STD E	F	PROB>P
INTERCEPT ASVABET ASVABCM ASVABCE	0.2686742 0.2374158 -0.3717212	0.1179 0.1197 0.1260	5.19 3.93 8.70	0.0236 0.0495 0.0035
	AT White	Women on	BADGUY	
REGRESSION ERRCR	236 36	70F 500 86.3025 67.6069	5.22	PROB>F 0.0001
	236 36	OF 500 86.3025	F	PROB>F 0.0001 PROB>F

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